# Scanner Data on Durable Goods: Market Dynamics and Hedonic Time Dummy Price Indexes 

Discussion paper 04011

## Explanation of symbols

- $\quad=$ data not available
* = provisional figure
$x \quad=$ publication prohibited (confidential figure)
- $\quad=$ nil or less than half of unit concerned
- $\quad=$ (between two figures) inclusive
$0(0,0) \quad=$ less than half of unit concerned
blank = not applicable
2003-2004 = 2003 to 2004 inclusive
2003/2004 = average of 2003 up to and including 2004
2003/'04 = crop year, financial year, school year etc. beginning in 2003 and ending in 2004
Due to rounding, some totals may not correspond with the sum of the separate figures.


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#### Abstract

Summary: The treatment of durable goods and particularly the way quality adjustments are dealt with, is an important issue in the compilation of a consumer price index. This paper has two aims: first, to describe the consumer market for durables in the Netherlands and, second, to investigate the potential of hedonic regression for quality adjustments. To this end we utilize scanner data on televisions, refrigerators, washing machines and personal computers that represent the total population of sales. It is shown that the markets for these durables have a dynamic character: each period many new models appear and many obsolete models disappear. The market dynamics have implications for statistical agencies: samples should be updated regularly to remain representative, thus increasing the number of quality adjustments. This is where hedonics might help. We not only estimate hedonic models but estimate as well hedonic price indexes, using the time dummy variable method. To assess the various indexes and the official Dutch CPI, we take the so-called generalised Törnqvist index as our reference or benchmark index. Using a decomposition thereof, we show that for televisions, refrigerators and washing machines the matched-model Törnqvist index is a good approximation of the reference index, despite the dynamic market structure. Due to the high expenditure share of new and disappearing computer models the matched-model Törnqvist for PCs exhibits an increasing difference with respect to the reference index.


Keywords: consumer price index, hedonic regression, scanner data.

## 1. Introduction

In 1996 the Boskin Commission published their report (Boskin et al., 1996) which criticised the way the Bureau of Labor Statistics compiled the US Consumer Price Index (CPI). This criticism was directed to several aspects of the CPI, among which the way quality changes were handled was one of the most important. According to the Boskin Commission the US CPI overestimated inflation by 0,6 percentage points on an annual basis, due to insufficient corrections in case of quality changes. These findings were reason for the Dutch Advisory Committee on Price Statistics to encourage Statistics Netherlands in starting a research project to improve the Dutch CPI. This project started late 2000. Topics in this research were an annual revision of the base year, the introduction of a geometric average or a CES price index and the application of hedonic techniques when correcting for quality changes. See De Haan (2004) for an overview of the current state of affairs in this research project.

Hedonic regression is advocated in the new manual on the CPI (ILO, 2004) and it is the main subject in this paper. To perform research into this area Statistics Netherlands has
bought scanner data for televisions, refrigerators, washing machines and personal computers from market research company GfK-Benelux. Practically, these data cover the population of sales during the period 1999-2001. Average prices (unit values) and quantities are available at a detailed level. For each durable an extensive list of technical characteristics is known as well. This enables us to develop hedonic models based on the population of sales and to calculate all kinds of population-based hedonic indexes. The presence of quantities makes it possible to compare hedonic indexes calculated with different types of weights. The importance of weighting has been stressed in recent years by Silver (2002) and Diewert (2003).

The quantities also give insight in the way sales are distributed over the year and in the relative importance of brands and models. It gives as well an indication on the length of the lifetime of a specific model of a durable and the pace by which substitution takes place. For a statistical office this information gives insight in the frequency at which samples have to be updated.

Section 2 starts with a description of the consumer market for the four durables that are discussed in this paper. Section 3 describes the development and estimation of hedonic models. An important aspect is the stability of the model parameters. Both weighted and unweighted estimation procedures are applied. In this section computers are not included as they are dealt with extensively in Van Mulligen (2003). The results will be used in section 4 to estimate various hedonic time dummy indexes based on different types of weights. The central question is whether hedonic quality adjustments are necessary or a matched-model approach suffices. To answer this question we use a 'generalised' Törnqvist index as a benchmark index. This time dummy index coincides with the Törnqvist index in case all items of two consecutive periods can be matched. A decomposition of this index is used to evaluate whether for televisions, refrigerators, washing machines and PCs the conditions are satisfied under which a matched-model Törnqvist index approximates the generalised version. Section 5 concludes.

## 2. Household durables, 1999-2001

### 2.1 The data

The scanner data refer to 18 bimonthly periods covering the years 1999-2001, except for PCs for which monthly data are available. The data set contains, per type of outlet, for each model sold unit values, quantities and a large set of characteristics. For the most important outlet types the data cover the whole population; all outlets belonging to those types are included. For the remaining outlet types the data are based on a sample and GfK subsequently raised the sample data to obtain population estimates. This means that the unit values and quantities reflect the population of sales in the Netherlands during the years 1999-2001.

The data received were subjected to a (limited) number of plausibility and consistency checks. It appeared that not all prices were presented in the same currency; some prices
in Dutch guilders had to be converted to euros. A limited number of records were deleted because prices were highly implausible. Inconsistencies between characteristics were corrected where possible, sometimes using information from catalogues and the Internet. Generally spoken, the data were of good quality.

### 2.2 Market of durables 1999-2001

During the three years 1999-2001 a huge number of durables has been sold as is shown in table 1. According to GfK, two outlet types mainly sell PCs for business purposes. In our analysis, we classified all sales of those two types as business sales. They count for $62 \%$ of total sales. In the following we restrict ourselves to private sales only.

Table 1. Number of durables sold during 1999-2001

| Type of durable | 1999 | 2000 | 2001 | $1999-2001$ |
| :--- | :---: | :---: | :---: | :---: |
| Televisions | 997310 | 1046023 | 965254 | 3008587 |
| Refrigerators | 547909 | 569770 | 565752 | 1638431 |
| - freestanding | 382207 | 401596 | 402225 | 1186028 |
| - built in | 165702 | 168174 | 163527 | 497403 |
| Washing machines | 458100 | 490285 | 504466 | 1452851 |
| Computers | 1039052 | 1011953 | 919634 | 2970639 |
| - business use | 668116 | 606104 | 570038 | 1844258 |
| - private use | 370936 | 405849 | 649596 | 1126381 |

Figure 1 shows the bimonthly sales during 1999-2001. A bit surprisingly, all durables exhibit a clear seasonal pattern. Televisions, computers (private sales only) and washing machines display a comparable pattern. All three have a top in their sales towards the end of the year. Mid 2000 shows an exception for televisions with an extra top halfway the year. This can be attributed to the European Championships Football, which were organised in Belgium and the Netherlands that year and the Olympic Games of September. Traditionally, big sport events stimulate the sales of televisions. For refrigerators the yearly peak is around August.


Figure 1. Durables, quantities sold during 1999-2001
As mentioned above, sales information is available for a number of outlet types. Table 2 shows their relative importance. For televisions, refrigerators and washing machines
chain stores and buying combinations are the most important outlet types. For televisions, independents also have a considerable market share. For refrigerators a separate type of outlet, kitchen retail, does exist. In most cases these outlets can be classified as chain stores. The majority of computer sales is realised in specialised stores. Note that freestanding and built-in refrigerators have different distributions. Built-in refrigerators are typically sold in kitchen retail stores, whereas the market share of this outlet type is negligible with regard to freestanding refrigerators.

Table 2. Relative importance of types of outlets by durable, based on units sold

| Type of outlet | Televisions | Refrigerators |  |  | Washing machines | Computers |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Freestanding | Built-in | Total |  |  |
| Buying ${ }^{\text {a }}$ | 28.8\% | 31.8\% | 6.6\% | 24.4\% | 35.2\% | 3.7\% |
| Chain stores | 49.3\% | 55.9\% | 2.0\% | 39.9\% | 52.7\% | 23.7\% |
| Depmoh ${ }^{\text {b) }}$ | 3.3\% | 3.0\% | 0.0\% | 2.1\% | 3.5\% | 4.9\% |
| Independents | 16.3\% | 7.8\% | 1.3\% | 5.8\% | 8.6\% | 5.4\% |
| Kitchen retail | - | 1.5\% | 90.1\% | 27.7\% | - | - |
| Photo retail | 2.3\% | - | - | - | - | 9.3\% |
| Computer stores | - | - | - | - | - | 53.0\% |
| Total sales | 100\% | 100\% | 100\% | 100\% | 100\% | 100\% |
| ${ }^{\text {a) }}$ Buying: Buying combinations <br> ${ }^{\text {b) }}$ Depmoh: Department stores and mail order hou |  |  |  |  |  |  |

Table 3 shows the number of brands available and the market share of the ten most important ones. A surprisingly large number of brands is active in the market of durables. However, based on expenditures most of these are unimportant or even negligible. The market of durables is dominated by a relatively small number of brands.

Table 3. Number of brands and relative importance of top 10 , based on expenditures

| Type of durable | Total number of brands | Market share of top 10 of brands |
| :--- | :---: | :---: |
| Televisions | 77 | $94 \%$ |
| Refrigerators | 74 | $78 \%$ |
| Washing machines | 45 | $90 \%$ |
| Computers | 33 | $95 \%^{\text {a) }}$ |

${ }^{\text {a) }}$ Included is brand 'clone' having $26 \%$ of the market

An aspect which is important for price index statisticians is the life cycle of an item. Price indexes are preferably compiled by comparing identical items (compare like with like). When items disappear from the market, a comparison is made between the 'old' disappearing item and a newly selected one. In most cases it will be necessary to adjust the price difference between the two for a difference in quality. These quality adjustments contain a serious risk for biases as is indicated by Boskin et al. (1996). So, the longer items are available (with significant sales) the easier the life for the price statistician. Figures 2 and 3 provide an indication of the period items ${ }^{1}$ are available on

[^0]the market. Figure 2 shows which part of the $199902^{2}$ population of items is still sold in period $t$. Figure 3 compares consecutive periods. For televisions, refrigerators and washing machines these are bimonthly periods, for computers the comparison is made on monthly data.


Figure 2. Number of items sold in both 199902 and t as percentage of number of items sold in 199902 (for computers 199901 and $t$ )


Figure 3. Number of items sold in both $t-1$ and $t$ as percentage of number of items of $t-1$ (bimonthly for televisions, refrigerators and washing machines, monthly for computers)

For televisions, refrigerators and washing machines the match is fairly stable. In each period between $74 \%$ and $80 \%$ of the items of the previous period is still available. For refrigerators, the fraction of matched items is on average a bit lower than for televisions and washing machines. So, each period between $20 \%$ and $25 \%$ of the items disappear. After three years only $20 \%$ of the items sold in 199902 are still available. For computers the average lifetime is much shorter. Each month nearly $33 \%$ of the items of the

[^1]previous month disappear. In a year time less than $10 \%$ of the initial population has remained. In fact, it is surprising that three computer models sold in 199902 are still on the market in 200112. These figures make clear that the CPI samples will have to be updated regularly to remain representative for the market as a whole. Appendix 1 illustrates whether Statistics Netherlands has been successful in maintaining representative samples.

Note that figures 2 and 3 may slightly overestimate the true attrition rates. For example, in figure 3 it is assumed that an item has disappeared from the market in period $t$ when the sales in that period have dropped to zero. It does happen, however, that in period $t+1$ (or later) the item is sold again. We have considered these sales as the introduction of a new item.

Combining the information about sales and life cycles does change the picture a bit. For all durables the total number of items can be divided in a relatively small number that remains on the market for a long time, the major sellers, and a majority of items having a short lifetime and modest sales. As illustration, table 4 shows the number of months with positive sales for refrigerators and computers. Note that there may be months with zero sales in between, so the total period for which an item is on the market is longer in most cases. For refrigerators, the table shows that only $6.2 \%$ of all items stay on the market more than 2.5 years, but these items are responsible for $37.3 \%$ of total expenditures. On the other hand, almost half of the items (46.1\%) is on the market less than six months, but these count for only $5.4 \%$ of total expenditures. For computers, the average lifetime is clearly shorter, $79 \%$ of the items stay on the market less than half a year, but only one third of total expenditures is realised by them. Slightly over $5 \%$ of the items has a lifetime longer than one year, but these computers are responsible for $30 \%$ of total expenditures.

Table 4. Length of sales period for refrigerators and computers

| No of months with <br> positive sales | Share of items |  |  | Part of total expenditures |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  | Refrigerators | Computers |  | Refrigerators | Computers |
| $1-6$ | $46.1 \%$ | $79.0 \%$ |  | $5.4 \%$ | $34.4 \%$ |
| $7-12$ | $19.6 \%$ | $15.8 \%$ |  | $9.8 \%$ | $35.1 \%$ |
| $13-18$ | $11.8 \%$ |  |  | $10.0 \%$ | $21.6 \%$ |
| $19-24$ | $9.6 \%$ | $0.7 \%$ |  | $18.1 \%$ | $7.3 \%$ |
| $25-30$ | $6.8 \%$ | $0.2 \%$ |  | $19.5 \%$ | $1.5 \%$ |
| $31-36$ | $6.2 \%$ | $0.0 \%$ |  |  | $37.3 \%$ |

Table 4 also may give a slightly distorted picture as items with sales only in the first or the last period of 1999-2001 actually can be sold much longer. In our table such an item is treated as an item with a very short lifetime. So, the presented information on lifetime and percentages of matched items has to be handled with some care. Nevertheless, it provides useful insight in the dynamic structure of the market of durables.

Figure 4 shows that over the three-year period the sales-weighted average prices (unit values) increase considerably. Refrigerators, washing machines and computers (private
use) exhibit an increase of around $10 \%$, whereas for televisions the increase is nearly twice as high (19\%). However, this increase has nothing to do with inflation but is determined by an increase in quality of the durables involved. Quality adjusted price indexes decline substantially, as will be shown in section 4.


Figure 4 . Durables, unit values $(€)$ during 1999-2001

## 3. Hedonic models

### 3.1 Introduction

A hedonic function describes the relation between the price of an item and a number of performance characteristics. Well-known functional specifications are the linear model, the semi-logarithmic (log-linear) model and the doublelog model. In the latter model not only the logarithm of the price but also the logarithm of the continuous explanatory variables is taken. The linear hedonic model can be expressed as

$$
\begin{equation*}
p_{i}^{t}=\alpha^{t}+\sum_{k=1}^{K} \beta_{k}^{t} X_{i k}+\sum_{j=1}^{J} \gamma_{j}^{t} D_{i j}+\varepsilon_{i}^{t} \tag{1}
\end{equation*}
$$

where $p_{i}^{t}$ denotes the price of item $i(i=1, \ldots, n)$ in period $t, X_{i k}$ its $k$-th quantitative characteristic $(k=1, . ., K), D_{i j}$ its $j$-th qualitative characteristic (i.e. a dummy variable, $j=1, \ldots, J), \beta_{k}^{t}$ and $\gamma_{j}^{t}$ the corresponding parameters and $\varepsilon_{i}^{t}$ an independently distributed error term with expected value of 0 and constant variance $\sigma^{2}$. Note that the parameters are modelled time-dependent as there is no a priori reason to assume they are constant over time.

Research on hedonic modelling for computers, based on the GfK-data, has been done by Van Mulligen (2003). The reader is referred to his chapter 4 for a detailed explanation of the models used. We will use his findings in our section 4 where we apply his loglinear model to calculate time dummy indexes for PCs. Dummy variables for brand name and outlet type are included as explanatory variables in the hedonic models for
computers. The estimated hedonic models for televisions, refrigerators and washing machines also include dummies for brand name and type of outlet.

We deleted built-in refrigerators from our database as expenditures on those goods are defined as gross capital formation by households in the system of National Accounts and in the Dutch CPI as well.

A subsequent data reduction has been carried out by deleting brands and items with a negligible market share. There are many brands having hardly any sales as was shown in table 3. Moreover, many items are available very temporarily, sometimes only one or two months as was illustrated in table 4 . To avoid disturbances due to incidental sales these transactions have been excluded from our analysis. Silver (2002), who discusses the use of weights in hedonic regressions, advises to delete so-called influential outliers, i.e. items with a small weight that have a large effect on the hedonic coefficients. Similar advice is given by Deltas and Zacharias (2004) who concluded to a bias in a high-frequency matched-model index due to models that survive for only a very small number of periods. During the process of estimating the hedonic models we encountered items with extreme (positive or negative) residuals. Further investigation learned that quite often these durables were of a specific design. Since we had no characteristics available relating to design features we deleted these items to prevent missing variable bias. In a few cases the extreme residuals related to very expensive or very cheap items. The most extreme cases have been deleted as well. Nevertheless, an impressive amount of data was left for our analysis. Table 5 gives an impression.

Table 5. Number of observations (unit values per item per period)

| Type of durable | No of observations | \% of total expenditures |
| :--- | :---: | :---: |
| Televisions | 24403 | $97.8 \%$ |
| Refrigerators | 18808 | $95.8 \%$ |
| Washing machines | 18478 | $96.4 \%$ |

We tested which of the three above-mentioned functional forms fitted best to the available data. The following test has been used, based on Box-Cox transformations (Lim and McKenzie, 2002).

For example, under the null hypothesis
$H_{0}$ : the linear and the log-linear model are empirically equivalent
the statistic $l=\frac{n}{2}\left|\ln \left(\frac{R S S_{\text {lin }} / \bar{y}^{2}}{R S S_{\log l i n}}\right)\right|$, where $\bar{y}^{2}=\exp \left(\frac{1}{n} \sum_{i=1}^{n} \ln p_{i}^{t}\right)$ and $R S S$ denotes the residual sum of squares, has a $\chi_{(1)}^{2}$ distribution.

This statistic has been used to evaluate all three possible combinations. The factor $\bar{y}^{2}$ is needed to compare the residuals of the linear model with those of the logarithmic
models. It should be omitted when testing the log-linear versus the doublelog model. In performing these tests no weights were used in the regression (Ordinary Least Squares, OLS). Regressions for periods 199902, 200008 and 200112 indicated that the log-linear and the doublelog model fitted significantly better to the data than the linear model did. Of these two, the log-linear form performed slightly better. The difference was marginal in most cases, which is probably due to the fact that nearly all explanatory variables are dummies. We have chosen the log-linear form in our subsequent analysis for its simpler interpretation and wide use in other empirical studies.

### 3.2 Hedonic regressions (OLS)

### 3.2.1 Results

For televisions, refrigerators and washing machines OLS-regressions have been run on the data of each of the 18 bimonthly periods. The explanatory variables used are listed in appendices 2-4. Table 6 contains some summarising statistics: the goodness of fit (adjusted $R^{2}$ ), the standard error of the estimate $(S E E)$, the number of items and the number of explanatory variables (including the intercept). The average, minimum and maximum value are based on their values for the 18 bimonthly regressions separately. Consequently, the standard deviation in table 6 gives an impression of the stability of the indicator over the three-year period.

Table 6. Summary results of OLS-regression on 18 bimonthly periods

| Type of durable |  | Average | St. Deviation | Minimum | Maximum |
| :--- | :--- | :---: | :---: | :---: | :---: |
| Televisions | Adjusted $\mathrm{R}^{2}$ | 0.963 | 0.004 | 0.957 | 0.973 |
|  | SEE | 0.144 | 0.013 | 0.117 | 0.159 |
|  | No of items | 1356 | 121 | 1095 | 1485 |
|  | No of variables | 63 | 3 | 57 | 66 |
| Refrigerators | Adjusted $\mathrm{R}^{2}$ | 0.886 | 0.024 | 0.851 | 0.930 |
|  | SEE | 0.180 | 0.028 | 0.130 | 0.219 |
|  | No of items | 1045 | 132 | 762 | 1268 |
|  | No of variables | 66 | 3 | 62 | 70 |
| Washing machines | Adjusted $\mathrm{R}^{2}$ | 0.877 | 0.022 | 0.847 | 0.932 |
|  | SEE | 0.131 | 0.011 | 0.098 | 0.144 |
|  | No of items | 898 | 74 | 757 | 998 |
|  | No of variables | 48 | 2 | 44 | 52 |

Figure 5 depicts the adjusted $R^{2}$ for all periods. Generally spoken, the goodness of fit decreases over time. The reason for this is not clear. Possibly, certain characteristics with an increasing influence on the price are missing in the GfK-database. The figure also shows that the fit is significantly better for televisions than for the other two durables. A summary of the regression results including the average value of the coefficients and their standard errors taken over all time periods is given in appendices $6-8$. For confidentiality reasons, brand names are anonymised. Detailed results (apart from brand names) are available from the author.


Figure 5. Development of adjusted $R^{2}$ (OLS)
For televisions 29 technical characteristics are incorporated into the hedonic model. The coefficients of 10 of these differ significantly from zero at the $5 \%$-level in all 18 periods and their signs are according to a priori expectations. Based on the (average) value of the coefficients the availability of a DVD-player is the most important price-increasing characteristic. This feature, however, is only available from 200103 onwards. Characteristics having a major price-increasing influence during the whole period are a flat screen and a 100 Hz frequency. The absence of teletext or stereo sound lowers the price, other things being equal.

For washing machines the coefficients of 4 out of the 15 technical characteristics differ significantly from zero in all 18 periods. Here as well, their signs are in accordance with a priori expectations. The availability of a separate spin-dryer has the largest (negative) coefficient.

For refrigerators the coefficients of 6 out of 20 technical characteristics are significantly different from zero in all 18 periods and their signs accord with a priori expectations. Having three doors raises the price sharply but only very few refrigerators have this feature. Fridge-freezers are on average $39 \%$ more expensive than refrigerators, with or without a freezing compartment.

In addition to the technical characteristics the brand name appears to be an important factor in explaining the price of a durable. The tables in the appendices 6-8 indicate that many brand names have significant coefficients. The effect of the brand name on the price of a durable is illustrated in table 7 that shows the price differences between a certain brand name and the brand name of the base item (benchmark item) in the regression model. For each durable the ten brands with the highest market share are listed. Because the brand of the base model always belongs to this top10, the table contains nine brands only.

Table 7. Top10 of brands; price level compared with base-brand ${ }^{\text {a) }}$

| Televisions |  |  | Refrigerators |  |  | Washing machines |  |
| :---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
|  | Brand | Price diff. |  | Brand | Price diff. |  |  |
| Brand08 | $+283 \%$ |  | Brand01 | $-2 \%$ |  |  | Brand07 |
| Brand21 | $+9 \%$ |  | Brand08 | $-14 \%$ |  |  | $-10 \%$ |
| Brand. |  |  | Brand11 | $-1 \%$ |  |  |  |
| Brand27 | $+7 \%$ |  | Brand10 | $+1 \%$ |  | Brand12 | $-8 \%$ |
| Brand31 | $+61 \%$ |  | Brand53 | $-18 \%$ |  | Brand28 | $-38 \%$ |
| Brand41 | $+7 \%$ |  | Brand56 | $+11 \%$ |  | Brand30 | $-12 \%$ |
| Brand42 | $+9 \%$ |  | Brand60 | $+30 \%$ |  | Brand31 | $+35 \%$ |
| Brand53 | $-6 \%$ |  | Brand67 | $-13 \%$ |  | Brand38 | $-1 \%$ |
| Brand59 | $-8 \%$ |  | Brand74 | $-1 \%$ |  | Brand42 | $-23 \%$ |
| Brand62 | $+17 \%$ |  | Brand76 | $+83 \%$ |  | Brand44 | $-21 \%$ |

a) Base-brand itself also belongs to top 10

Table 7 shows that brand effects can be substantial. The most extreme case is brand 08 for televisions which suggests that consumers are willing to pay nearly three times as much for a television set of this brand than for a model with the same characteristics of the reference brand. This behaviour does not seem realistic. The extreme value is probably, at least partly, due to the ambiguous nature of the coefficient of the brand name. It is well-known that brands can have a certain reputation for which consumers are willing to pay. Quite often, this reputation is historically based and cannot explicitly be attributed to specific features. This causes brand names having a (positive or negative) influence on the price of a durable. Another, and maybe more important, problem is that not all price determining factors are known or available. A quality difference between two models of a durable good can be difficult to quantify. Examples are the quality of the components or the appearance of a durable, i.e. the way it is designed. A specific styling can be very appealing to a certain group of consumers who are willing to pay a considerable bonus for it. When no specific variables are available to describe these features the price effect will be included in the coefficient of the brand name. In that sense the brand name is a kind of garbage bin of the effect of all characteristics which have not been included in the regression model.

OLS-regression assumes that the errors are independently and identically distributed with expectation zero and constant variance (homoscedasticity). No formal tests to check these conditions have been applied. We did, however, analyse graphs of residuals. Those graphs did not show strong signs of heteroscedasticity. We refer to the detailed reports for some figures (Van der Grient, 2003a, 2003b and Oei, 2003).

### 3.2.2 Stability of coefficients with OLS regression

The hedonic technique can be used in several ways when correcting for quality changes in the CPI. Triplett (2003) gives an overview. One of the methods is the time dummy variable method, which will be further explored in section 4 . An important condition which must be met when applying this variant is the constancy (stability) of the regression coefficients over time, at least in two consecutive periods. A first impression of this kind of stability is presented in figures 6-8 which show the OLS coefficient
estimates for the most important (technical) characteristics of each durable. All these coefficients differ significantly from zero at the $5 \%$-level in all 18 periods. In general, instability does not seem to pose a serious problem.


Figure 6. Televisions: OLS regression coefficients


Figure 7. Refrigerators: OLS regression coefficients


Figure 8. Washing machines: OLS regression coefficients

We will now perform a formal test. Special attention will be paid to the length of the period over which coefficients can be considered constant. The constancy is tested using a Chow-test statistic in which the residual sum of squares of separate regressions in two periods is compared with the residual sum of squares of the regression on the data of both periods. In this pooled regression a time dummy variable has been added to correct for a general price change between the two periods involved. The log-linear hedonic models for the unpooled and pooled data read

$$
\begin{equation*}
\ln \left(p_{i}^{t}\right)=\alpha^{t}+\sum_{k=1}^{K} \beta_{k}^{t} X_{i k}+\sum_{j=1}^{J} \gamma_{j}^{t} D_{i j}+\varepsilon_{i}^{t} \quad \quad(\mathrm{t}=0,1) \tag{2}
\end{equation*}
$$

and

$$
\begin{equation*}
\ln \left(p_{i}^{t}\right)=\alpha+\delta T_{i}^{t}+\sum_{k=1}^{K} \beta_{k} X_{i k}+\sum_{j=1}^{J} \gamma_{j} D_{i j}+\varepsilon_{i}^{t} \quad(\mathrm{t}=0,1) \tag{3}
\end{equation*}
$$

respectively, where $T_{i}^{t}$ in (3) is a dummy variable that takes on the value of 1 if the observation comes from period 1 ( 0 otherwise).

Under the null-hypothesis $H_{0}: \alpha_{t}=\alpha_{t-1}, \beta_{k, t}=\beta_{k, t-1}$ for $k=1$..K
The test-statistic $F_{[K, n-2 K]}=\frac{\left[R S S_{0+1}-R S S_{0}-R S S_{1}\right] / K}{\left[R S S_{0}+R S S_{1}\right] /(n-2 K)}$
has an F-distribution with $K$ and $n-2 K$ degrees of freedom where $n$ is the total number of observations in periods 0 and 1 combined. Rejection of the null-hypothesis means that one or more of the coefficients are not constant. Table 8 shows the results of this test. The coefficients appear to be fairly stable. Over a period of four months (a comparison between $t$ and $t+2$ ) constancy of coefficients is never rejected. Over the period of half a year the hypothesis on constancy has to be rejected three times for televisions and only once for washing machines. In chapter 4 only adjacent periods will be combined. Apparently this may be done without violating the assumption of stability.

Table 8. Test on stability of coefficients (OLS) at a 5\% significance level

| Periods <br> compared |  | Number of combined periods with stable and unstable coefficients |  |  |
| :---: | :--- | :---: | :---: | :---: |
|  | Stable | Televisions | Refrigerators | Washing machines |
|  | Not stable | 17 | 17 | 17 |
| $\mathrm{t}, \mathrm{t}+2$ | Stable | 0 | 0 | 0 |
|  | Not stable | 16 | 16 | 16 |
| $\mathrm{t}, \mathrm{t}+3$ | Stable | 0 | 0 | 0 |
|  | Not stable | 12 | 15 | 14 |
| $\mathrm{t}, \mathrm{t}+4$ | Stable | 3 | 0 | 1 |
|  | Not stable | 8 | 11 | 10 |
| t, $\mathrm{t}+5$ | Stable | 6 | 3 | 4 |
|  | Not stable | 7 | 9 | 8 |
| t, $\mathrm{t}+6$ | Stable | 6 | 4 | 5 |
|  | Not stable | 3 | 7 | 2 |

a) $t=199902, \ldots, 200110$

### 3.3 Hedonic regressions (WLS)

### 3.3.1 Results

There may be two reasons to prefer weighted least squares regression (WLS) over unweighted regression (OLS). First, WLS can correct for heteroscedasticity. Second, items with high expenditures or sales should have a bigger impact than low-expenditure ones. The last point is particularly important for the time dummy variable method where the estimated hedonic model directly leads to a price index (see for example Silver, 2002 and Diewert, 2003)

Although the OLS-regression did not point to strong heteroscedasticity, we ran (expenditure) weighted regressions to see the effect weighting has on the coefficients. It was assumed that the log-linear functional form is the optimal one here as well although this was not tested formally. In table 9 summarising statistics are given.

Table 9. Summary results of WLS-regression (expenditure-weighted) in 18 periods

| Type of Durable |  | Average | St. Deviation | Minimum | Maximum |
| :--- | :--- | :---: | :---: | :---: | :---: |
| Televisions | Adjusted $\mathrm{R}^{2}$ | 0.969 | 0.005 | 0.962 | 0.979 |
|  | SEE | 0.107 | 0.007 | 0.095 | 0.115 |
|  | No of cases | 1356 | 121 | 1095 | 1485 |
| Refrigerators | No of variables | 63 | 3 | 57 | 66 |
|  | Adjusted R |  | 0.017 | 0.902 | 0.954 |
|  | SEE | 0.928 | 0.015 | 0.123 | 0.179 |
|  | No of cases | 0.145 | 1045 | 132 | 762 |
| Washing machines | No of variables | 66 | 3 | 62 | 1268 |
|  | Adjusted R |  | 70 |  |  |
|  | SEE | 0.939 | 0.009 | 0.921 | 0.965 |
|  | No of cases | 0.096 | 0.007 | 0.079 | 0.106 |
|  | No of variables | 898 | 74 | 757 | 998 |

We compare the WLS-results given in table 9 with the corresponding OLS-results of table 6. For televisions the fit hardly improves. For the other two durables, on the other hand, the adjusted $R^{2}$ improves considerably. Moreover, the adjusted $R^{2}$ from the weighted regression appears to be much more stable over time which follows from a comparison between the standard deviations in tables 6 and 9 . The standard error of the estimate (SEE) also indicates that weighted regression is to be preferred in terms of goodness of fit. The adjusted $R^{2}$ decreases over time here as well, but less sharply as was the case with OLS. Graphically this is shown in figure 9. A comparison between figures 5 and 9 illuminates the remarks made above.


Figure 9. Development of adjusted $R^{2}$ (WLS, expenditure-weighted)
The 'most important' characteristics found with OLS are also the most important ones using WLS. The number of variables whose coefficients differ significantly from zero in all 18 periods is a bit lower, however. In most cases this is due to the situation in one particular period. Sales and so the relative importance of items sometimes change dramatically between two time periods and this affects the WLS-estimates. Summary results for all the explanatory variables are given in appendices $10-12$. Details are available from the author.

### 3.3.2 Stability of coefficients with WLS regression

As the adjusted $R^{2}$ is more stable using WLS, the question arises whether the same holds for the coefficients of the most important characteristics. Figures 10-12 depict the WLS-coefficients of the same technical characteristics as were presented in figures 6-8 for the OLS-case. Surprisingly, the WLS-coefficients are slightly more volatile than their OLS-counterparts. Furthermore, four coefficients (two for refrigerators and two for washing machines) are no longer significantly different from zero in all 18 periods. So generally spoken, the coefficients from the weighted regressions are a bit less stable than those from the unweighted regressions. The standard deviations of the coefficients for the 18 periods, presented in appendices $10-12$ show that on average the WLScoefficients have a somewhat greater dispersion than the OLS-coefficients.


Figure 10. Televisions: WLS regression coefficients


Figure 11. Refrigerators: WLS regression coefficients


Figure 12. Washing machines: WLS regression coefficients
The stability of the WLS-coefficients has also formally been tested in the same way as was done for the OLS-coefficients. In the pooled regressions, expenditure shares have been used instead of the expenditures themselves. Diewert (2003) points out that expenditure share weights should be used as opposed to expenditures to avoid inflation
increasing period 1 value weights resulting in possible heteroscedastic residuals. The results of the test are shown in table 10 . A comparison of this table with table 8 confirms the impression of less stable WLS-coefficients. Even between two consecutive periods the assumption of constant coefficients is sometimes rejected. For longer intervals the instability increases sharply.

Table 10. Test on stability of coefficients (WLS) at a 5\% significance level

| Periods compared |  | Number of combined periods with stable and unstable coefficients |  |  |
| :---: | :---: | :---: | :---: | :---: |
|  |  | Televisions | Refrigerators | Washing machines |
| $\mathrm{t}, \mathrm{t}+1^{\text {a) }}$ | Stable | 16 | 14 | 17 |
|  | Not stable | 1 | 3 | 0 |
| t, t+2 | Stable | 8 | 8 | 12 |
|  | Not stable | 8 | 8 | 4 |
| t, t+3 | Stable | 6 | 7 | 4 |
|  | Not stable | 9 | 8 | 11 |
| $\mathrm{t}, \mathrm{t}+4$ | Stable | 4 | 1 | 3 |
|  | Not stable | 10 | 13 | 11 |
| $\mathrm{t}, \mathrm{t}+5$ | Stable | 2 | 0 | 1 |
|  | Not stable | 11 | 13 | 12 |
| $t, t+6$ | Stable | 2 | 0 | 0 |
|  | Not stable | 10 | 12 | 12 |

a) $\mathrm{t}=199902, \ldots, 200110$

One of the main aims in this project was to develop hedonic models, based on the population of sales, for specific durable goods. When using the hedonic technique in CPI-practice, it could be advisable to estimate models on subpopulations that are more homogeneous than the total population. We tested this for refrigerators. An (expenditure shares) weighted regression was run on a subpopulation consisting of the ten most important brands, after a further reduction of items with low sales. The stability of coefficients increased considerably. The null-hypothesis of stability between adjacent periods was never rejected and only once for a comparison between periods $t$ and $t+2$.

## 4. Time dummy indexes for durables

### 4.1 Introduction

A well-known method (although not frequently used by statistical agencies) to estimate a hedonic price index is the time dummy method. Consider two periods, denoted by 0 and 1 . Recall the pooled version of the log-linear hedonic model already used in section 3.2.2
$\ln \left(p_{i}^{t}\right)=\alpha+\delta T_{i}^{t}+\sum_{k=1}^{K} \beta_{k} X_{i k}+\sum_{j=1}^{J} \gamma_{j} D_{i j}+\varepsilon_{i}^{t} \quad(\mathrm{t}=0,1)$

Note again that the $\beta_{k}$ 's and $\gamma_{j}$ 's are assumed constant over time. For televisions, refrigerators and washing machines this restriction does not seem a serious problem, for computers we have some doubts whether this condition is met. Due to an insufficient number of cases Van Mulligen (2003) did not run period-specific regressions on the computer data and did not test for stability.

Model (4) will be estimated, using weighted and unweighted regression, on the pooled data from both periods. The estimated or predicted price of $i$ is $\hat{p}_{i}^{t}$, the residual $u_{i}^{t}=\ln \left(p_{i}^{t}\right)-\ln \left(\hat{p}_{i}^{t}\right)=\ln \left(p_{i}^{t} / \hat{p}_{i}^{t}\right)$. The antilogarithm (exponent) of the time dummy coefficient $\hat{\delta}$ automatically produces a quality-adjusted price index. In the case of WLS we will apply three types of weighting, expenditure shares, quantities and a specific type of average weights. This average weight is based on Diewert (2003) who suggests using the average expenditure shares of both periods as weights for an item that has been sold in both periods. For new and disappearing items, which are by definition available in one period only, the expenditure shares relating to that period should serve as weights. If there happen to be no new or disappearing items, in which case there are only matched items, this WLS-estimator coincides with the Törnqvist index. According to Diewert the resulting WLS-estimator provides a generalisation of the (superlative) Törnqvist index ${ }^{3}$. We believe that his procedure, when applied to the whole population, offers a benchmark index that can be used to assess other indexes ${ }^{4}$. Using a specific decomposition of the generalised Törnqvist index we will investigate if a matchedmodel index can be considered to be a good approximation thereof.

### 4.2 Results

Figures 13-16 depict several time dummy indexes and the importance of weighting is clearly demonstrated. For televisions, refrigerators and washing machines, the OLS time dummy index approximates the generalised Törnqvist index, WLS(Gen. Törnqvist), surprisingly well, although both indexes seem to diverge at the end of the period studied. The use of time-specific expenditure shares as weights, on the other hand, results in a widening gap with the benchmark index. For computers there is a great similarity between both WLS time dummy indexes. The OLS index is slightly higher.

[^2]

Figure 13 Televisions, three time dummy indexes and the CPI


Figure 14. Refrigerators, three time dummy indexes and the CPI


Figure 15. Washing machines, three time dummy indexes and the CPI


Figure 16. Computers, three time dummy indexes and the CPI
To keep the figures surveyable, the quantity weighted time dummy indexes are not shown. They can be found in appendices 14-17 together with the indexes shown in the figures. For refrigerators, washing machines and computers the quantity weighted time dummy indexes are closer to the generalised Törnqvist than the expenditure-share weighted ones. Detailed regression results for televisions, refrigerators and washing machines are not presented, because the average coefficients of the pooled regressions only slightly differ from those of the period-specific regressions. This is the case for OLS- and for WLS-regression, using the same type of weight. Therefore only the summary results for computers are presented in appendices 9 and 13.

The CPI-figures, which are presented in figures 13-16 as well, diverge significantly from the benchmark index ${ }^{5}$. For televisions, refrigerators and washing machines the CPI seems to be upwardly biased. Surprisingly, the CPI for PCs is lower than all time dummy indexes. Another remarkable point is that the CPI and the expenditure-share weighted time dummy index converge sharply in the last quarter of 2001 for televisions, refrigerators and computers. In fact, both indexes coincide in the last period.

One should bear in mind that the CPI methodology and the data used differ in many respects from the time dummy indexes. The CPI is based on a relatively small sample (between 15 and 20 models for televisions and refrigerators and less than 10 models for washing machines) for which prices are collected on a monthly basis. It is a Laspeyrestype index; the (constant) weights reflect the 1995 market share of the brands involved. The number of explicit quality adjustments is limited; in practice a matched-model

[^3]approach is applied. The official price index for computers is essentially an unweighted matched model chain index. Prices of identical computers per retailer are compared. New computers are incorporated using the class mean method, so no explicit quality adjustments are made. See Van Mulligen (2003) for the details. Finally, the CPI is based on list prices, not on (average) transaction prices as are the time dummy indexes. So, there are several aspects that contribute to the difference between the CPI and the population based time dummy indexes. Our findings about the representativity of the CPI-sample, discussed in appendix 1, suggest that sampling aspects play an important role. Van Mulligen (2003) also concluded that for PCs "the sampling and weighting issues are more likely candidates for the widening gap between the matched-model (scanner data based) indices and the CPI" ${ }^{6}$.

### 4.3 Decomposition of the WLS time dummy price index

### 4.3.1 Introduction

Recently, Van der Grient and De Haan (2003) showed that for televisions the matchedmodel Törnqvist index is a very good approximation of the generalised Törnqvist index $\left(P_{G T}\right)$. Their analysis was based on a decomposition of $P_{G T}$. In this section a similar analysis will be performed for refrigerators, washing machines and computers. For comparison, the results for televisions will be repeated.
Based on hedonic model (4), Van der Grient and De Haan (2003) demonstrated that the resulting time dummy index can be written as

$$
\begin{equation*}
P_{G T}=\left(P_{M G L}\right)^{\frac{s_{M}^{0}}{s_{M}^{O}+s_{M}^{1}}}\left(P_{M G P}\right)^{\frac{s_{M}^{1}}{s_{M}^{0}+s_{M}^{1}}}\left[\exp \left(\bar{u}_{N}\right)\right]^{\frac{2\left(1-s_{M}^{1}\right)}{s_{M}^{0}+s_{M}^{1}}}\left[\exp \left(-\bar{u}_{D}\right)\right]^{\frac{2\left(1-s_{M}^{0}\right)}{s_{M}^{0}+s_{M}^{1}}}, \tag{5}
\end{equation*}
$$

where $s_{M}^{t}$ denotes the period $t$ expenditure share $(t=0,1)$ of the matched population. The expenditure shares weighted average residuals of new and disappearing items are denoted by $\bar{u}_{N}=\sum_{i \in U_{N}} s_{i}^{1} u_{i}^{1} / \sum_{i \in U_{N}} s_{i}^{1}$ and $\bar{u}_{D}=\sum_{i \in U_{D}} s_{i}^{0} u_{i}^{0} / \sum_{i \in U_{D}} s_{i}^{0}$ where $s_{i}^{0}$ and $s_{i}^{1}$ represent the expenditure shares of item $i$ in periods 0 and 1 respectively. The other part of (5) relates to the weighted geometric average of two matched-model price indexes: the geometric Laspeyres index $P_{M G L}=\prod_{i \in U_{M}}\left(p_{i}^{1} / p_{i}^{0}\right)^{s_{i M}^{0}}$ and the geometric Paasche index $P_{M G P}=\prod_{i \in U_{M}}\left(p_{i}^{1} / p_{i}^{0}\right)^{s_{i M}^{1}}$, with $s_{i M}^{t}=s_{i}^{t} / s_{M}^{t}$ being the share of item $i$ in the period $t$ expenditures on the set $U_{M}$ of all matched items. In equation (5) the impact of new and disappearing items on $P_{G T}$ is made explicit. Note that all factors depend on the expenditure share of new and disappearing items.

The matched-model Törnqvist index is defined as the unweighted geometric average of $P_{M G L}$ and $P_{M G P}$, i.e.

[^4]$P_{M T}=\left(P_{M G L} P_{M G P}\right)^{1 / 2}$
Comparing equations (5) and (6) shows that two assumptions are needed to ensure that the matched-model Törnqvist index coincides with the generalised one:

Assumption $i: s_{M}^{1}=s_{M}^{0}$.
Replacing $s_{M}^{1}$ by $s_{M}^{0}$ in decomposition (5) gives

$$
\begin{equation*}
P_{G T(i)}=P_{M T}\left[\exp \left(\bar{u}_{N}-\bar{u}_{D}\right)\right]^{\frac{1}{s_{M}^{0}}-1} . \tag{7}
\end{equation*}
$$

Assumption ii: $\bar{u}_{N}=\bar{u}_{D}$.
Under this assumption $P_{G T(i)}$ reduces to $P_{M T}$.
Note that it is not the variability of the quality-adjusted prices of unmatched items that matters. Rather it is systematic effects - giving rise to positive or negative average residuals - that matter. Whether ii holds depends on the prevailing market conditions and the pricing strategies followed by manufacturers when introducing new models into the market.

### 4.3.2 Results

In this section we will calculate the factors of decomposition of equation (5) for the different durables and analyse whether assumptions $i$ and $i i$ will hold.

## Factors of decomposition

Table 11 contains the results of the decomposition of $P_{G T}$ according to expression (5). $F(L / P)=\left(P_{M G L}\right)^{\frac{s_{M}^{0}}{s_{M}^{0}+s_{M}^{1}}}\left(P_{M G P}\right)^{\frac{s_{M}^{1}}{s_{M}^{1}+s_{M}^{1}}}$ is the weighted average of the matched-model Laspeyres and Paasche indexes and $F$ (res) denotes the remaining part of the right-hand side of (5). For televisions, refrigerators and washing machines the effect of $F($ res $)$ in $P_{G T}$ appears to be negligible. For computers on the other hand, the effect is substantial. In the following we will analyse these aspects in greater detail and explain why computers behave so differently.

Table 11. $P_{G T}$ and factors of decomposition according to (5)

| Period | Televisions |  |  | Refrigerators |  |  | Washing machines |  |  | Computers |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $P_{G T}$ | $F(L / P)$ | $F$ (res) | $P_{G T}$ | $F(L / P)$ | $F(r e s)$ | $P_{G T}$ | $F(L / P)$ | $F$ (res) | $P_{G T}$ | $F(L / P)$ | $F$ (res) |
| 199901 |  |  |  |  |  |  |  |  |  | 100.00 | 100.00 | 1.000 |
| 199902 | 100.00 | 100.00 | 1.000 | 100.00 | 100.00 | 1.000 | 100.00 | 100.00 | 1.000 | 96.28 | 94.77 | 1.016 |
| 199903 |  |  |  |  |  |  |  |  |  | 92.71 | 61.68 | 1.011 |
| 199904 | 97.26 | 96.95 | 1.003 | 99.01 | 99.09 | 0.999 | 98.03 | 97.76 | 1.001 | 88.92 | 89.12 | 0.998 |
| 199905 |  |  |  |  |  |  |  |  |  | 85.68 | 86.99 | 0.985 |
| 199906 | 94.97 | 94.87 | 1.001 | 97.71 | 98.05 | 0.997 | 98.03 | 97.55 | 1.005 | 82.41 | 83.99 | 0.981 |
| 199907 |  |  |  |  |  |  |  |  |  | 77.26 | 78.12 | 0.989 |
| 199908 | 93.24 | 93.24 | 1.000 | 97.37 | 97.67 | 0.997 | 98.20 | 97.93 | 1.003 | 70.80 | 72.48 | 0.977 |
| 199909 |  |  |  |  |  |  |  |  |  | 67.47 | 69.15 | 0.976 |
| 199910 | 90.86 | 91.10 | 0.997 | 97.41 | 97.25 | 1.002 | 97.76 | 97.54 | 1.002 | 67.87 | 69.53 | 0.976 |
| 199911 |  |  |  |  |  |  |  |  |  | 66.31 | 68.15 | 0.973 |
| 199912 | 89.62 | 89.79 | 0.998 | 96.75 | 96.77 | 1.000 | 97.09 | 96.95 | 1.001 | 64.88 | 67.48 | 0.961 |
| 200001 |  |  |  |  |  |  |  |  |  | 64.30 | 66.51 | 0.967 |


| Period | Televisions |  |  | Refrigerators |  |  | Washing machines |  |  | Computers |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $P_{G T}$ | $F(L / P)$ | $F$ (res) | $P_{G T}$ | $F(L / P)$ | $F$ (res) | $P_{G T}$ | $F(L / P)$ | $F$ (res) | $P_{G T}$ | $F(L / P)$ | $F$ (res) |
| 200002 | 88.43 | 88.72 | 0.997 | 96.12 | 96.24 | 0.999 | 96.55 | 96.19 | 1.004 | 62.88 | 66.16 | 0.950 |
| 200003 |  |  |  |  |  |  |  |  |  | 62.59 | 65.78 | 0.952 |
| 200004 | 87.81 | 87.63 | 1.002 | 96.02 | 95.60 | 1.004 | 96.07 | 95.66 | 1.004 | 61.67 | 65.09 | 0.947 |
| 200005 |  |  |  |  |  |  |  |  |  | 60.31 | 63.66 | 0.947 |
| 200006 | 87.16 | 87.31 | 0.998 | 95.21 | 94.63 | 1.006 | 95.24 | 94.80 | 1.005 | 56.56 | 60.30 | 0.938 |
| 200007 |  |  |  |  |  |  |  |  |  | 55.72 | 59.26 | 0.940 |
| 200008 | 86.44 | 86.62 | 0.998 | 94.29 | 93.75 | 1.006 | 93.82 | 93.64 | 1.002 | 54.37 | 58.34 | 0.932 |
| 200009 |  |  |  |  |  |  |  |  |  | 53.16 | 57.65 | 0.922 |
| 200010 | 85.31 | 85.76 | 0.995 | 94.04 | 93.70 | 1.004 | 9308 | 93.00 | 1.001 | 51.69 | 56.21 | 0.920 |
| 200011 |  |  |  |  |  |  |  |  |  | 50.63 | 55.51 | 0.912 |
| 200012 | 85.17 | 85.40 | 0.997 | 94.30 | 93.59 | 1.008 | 92.66 | 92.48 | 1.002 | 50.62 | 54.72 | 0.925 |
| 200101 |  |  |  |  |  |  |  |  |  | 50.22 | 54.37 | 0.924 |
| 200102 | 85.55 | 85.70 | 0.998 | 95.16 | 94.25 | 1.010 | 92.91 | 92.82 | 1.001 | 47.99 | 52.34 | 0.917 |
| 200103 |  |  |  |  |  |  |  |  |  | 44.81 | 50.14 | 0.894 |
| 200104 | 85.72 | 85.68 | 1.000 | 95.76 | 95.07 | 1.007 | 93.68 | 93.45 | 1.002 | 44.39 | 49.76 | 0.892 |
| 200105 |  |  |  |  |  |  |  |  |  | 43.24 | 48.63 | 0.889 |
| 200106 | 85.75 | 85.74 | 1.000 | 96.00 | 94.44 | 1.017 | 93.59 | 93.04 | 1.006 | 41.95 | 47.43 | 0.884 |
| 200107 |  |  |  |  |  |  |  |  |  | 41.61 | 45.96 | 0.905 |
| 200108 | 84.90 | 84.75 | 1.002 | 94.72 | 93.55 | 1.012 | 92.82 | 92.57 | 1.003 | 38.59 | 44.29 | 0.871 |
| 200109 |  |  |  |  |  |  |  |  |  | 35.79 | 41.11 | 0.871 |
| 200110 | 83.99 | 84.08 | 0.999 | 94.58 | 93.39 | 1.013 | 92.32 | 92.07 | 1.003 | 35.49 | 40.28 | 0.881 |
| 200111 |  |  |  |  |  |  |  |  |  | 33.21 | 39.52 | 0.840 |
| 200112 | 83.25 | 83.42 | 0.998 | 94.57 | 93.19 | 1.015 | 91.58 | 91.21 | 1.004 | 32.22 | 38.77 | 0.831 |

## Expenditure shares of matched items

Appendix 18 shows the expenditure shares of the matched items in both period $t-1$ and $t$ for all adjacent periods. In table 12 the average values are presented. For most durables a very small part of the monthly expenditures is due to disappearing items. For televisions, for instance, the average expenditure share of disappearing items is $2 \%$ on a bimonthly basis, say $1 \%$ on a monthly basis. For refrigerators, this fraction amounts to $3.5 \%$. The part of turnover due to newly appearing items is nearly twice as high but still small compared to turnover based on items that are sold in both periods. As could be expected, the fractions of new and disappearing items for computers are much higher. Note that the expenditure shares for computers in table 12 are on a monthly basis, so each month $15 \%$ of turnover is lost due to disappearing items while turnover grows with $22 \%$ due to new items.

Table 12. Expenditure shares of the matched items in period $\mathbf{t}-\mathbf{1}$ and period $\mathbf{t}$; average over 199901-200112 ${ }^{\text {a) }}$

|  | Televisions |  | Refrigerators |  | Washing machines |  | Computers |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $s_{M}^{t-1}$ | $s_{M}^{t}$ | $S_{M}^{t-1}$ | $s_{M}^{t}$ | $S_{M}^{t-1}$ | $s_{M}^{t}$ | $S_{M}^{t-1}$ | $s_{M}^{t}$ |
| Average | 0.98 | 0.96 | 0.93 | 0.90 | 0.96 | 0.93 | 0.85 | 0.78 |

a) Bimonthly periods for televisions, refrigerators and washing machines, monthly periods for computers

These figures present a completely different picture of the market dynamics than figures 2 and 3 in section 2 did. Those figures were based on numbers sold, while the shares of table 12 are based on expenditures. The difference between sales and expenditures was already illustrated in table 4.
$P_{G T(i)}$, the first approximation of $P_{G T}$, is based on assumption $i$ that the expenditure shares of the matched items does not change between periods $t-1$ and $t$. According to table 12 the differences between the two seem to be small on average, at least for three of the four durable goods. For all four durables the consequences of assumption $i$ is shown in table 13. The differences between $P_{G T}$ and $P_{G T(i)}$ are indeed small, even for computers. In fact, the differences are almost negligible compared with the decline in the indexes during the period studied.

Table 13. $P_{G T}$ and $P_{G T(i)}$, effect of assumption $\boldsymbol{i}$

| Period | Televisions |  | Refrigerators |  | Washing machines |  | Computers |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $P_{G T}$ | $P_{G T(i)}$ | $P_{G T}$ | $P_{G T(i)}$ | $P_{G T}$ | $P_{G T(i)}$ | $P_{G T}$ | $P_{G T(i)}$ |
| 199901 |  |  |  |  |  |  | 100.00 | 100.00 |
| 199902 | 100.00 | 100.00 | 100.00 | 100.00 | 100.00 | 100.00 | 96.28 | 96.02 |
| 199903 |  |  |  |  |  |  | 92.71 | 92.44 |
| 199904 | 97.26 | 97.01 | 99.01 | 99.02 | 98.03 | 98.01 | 88.92 | 88.66 |
| 199905 |  |  |  |  |  |  | 85.68 | 85.70 |
| 199906 | 94.97 | 94.73 | 97.71 | 97.81 | 98.03 | 97.87 | 82.41 | 82.42 |
| 199907 |  |  |  |  |  |  | 77.26 | 77.13 |
| 199908 | 93.24 | 92.99 | 97.37 | 97.51 | 98.20 | 98.18 | 70.80 | 70.68 |
| 199909 |  |  |  |  |  |  | 67.47 | 67.32 |
| 199910 | 90.86 | 90.75 | 97.41 | 97.49 | 97.76 | 97.73 | 67.87 | 67.64 |
| 199911 |  |  |  |  |  |  | 66.31 | 66.29 |
| 199912 | 89.62 | 89.44 | 96.75 | 96.86 | 97.09 | 97.15 | 64.88 | 64.95 |
| 200001 |  |  |  |  |  |  | 64.30 | 64.35 |
| 200002 | 88.43 | 88.24 | 96.12 | 96.23 | 96.55 | 96.50 | 62.88 | 63.24 |
| 200003 |  |  |  |  |  |  | 62.59 | 63.11 |
| 200004 | 87.81 | 87.61 | 96.02 | 96.02 | 96.07 | 96.00 | 61.67 | 62.18 |
| 200005 |  |  |  |  |  |  | 60.31 | 60.87 |
| 200006 | 87.16 | 87.07 | 95.21 | 95.15 | 95.24 | 95.18 | 56.56 | 57.21 |
| 200007 |  |  |  |  |  |  | 55.72 | 56.27 |
| 200008 | 86.44 | 86.34 | 94.29 | 94.24 | 93.82 | 93.86 | 54.37 | 55.09 |
| 200009 |  |  |  |  |  |  | 53.16 | 54.14 |
| 200010 | 85.31 | 85.27 | 94.04 | 94.02 | 9308 | 93.15 | 51.69 | 52.65 |
| 200011 |  |  |  |  |  |  | 50.63 | 51.80 |
| 200012 | 85.17 | 85.01 | 94.30 | 94.22 | 92.66 | 92.69 | 50.62 | 51.48 |
| 200101 |  |  |  |  |  |  | 50.22 | 51.07 |
| 200102 | 85.55 | 85.34 | 95.16 | 95.09 | 92.91 | 92.93 | 47.99 | 48.80 |
| 200103 |  |  |  |  |  |  | 44.81 | 45.68 |
| 200104 | 85.72 | 85.48 | 95.76 | 95.76 | 93.68 | 93.64 | 44.39 | 45.38 |
| 200105 |  |  |  |  |  |  | 43.24 | 44.13 |
| 200106 | 85.75 | 85.52 | 96.00 | 95.97 | 93.59 | 93.46 | 41.95 | 42.77 |
| 200107 |  |  |  |  |  |  | 41.61 | 42.71 |
| 200108 | 84.90 | 84.56 | 94.72 | 94.80 | 92.82 | 92.73 | 38.59 | 40.37 |
| 200109 |  |  |  |  |  |  | 35.79 | 37.39 |
| 200110 | 83.99 | 83.68 | 94.58 | 94.07 | 92.32 | 92.16 | 35.49 | 36.86 |
| 200111 |  |  |  |  |  |  | 33.21 | 34.85 |
| 200112 | 83.25 | 82.97 | 94.57 | 94.85 | 91.58 | 91.39 | 32.22 | 33.95 |

## Average residuals (weighted) of new and disappearing items

Manufacturers can follow different strategies when introducing new models into the market. They can use the opportunity to increase prices more than can be justified by the change in quality (so-called hidden price increases). Another strategy may be to price new models relatively low with the aim of enlarging the market share. Schultze and Mackie (2001) quote BLS-research (Moulton and Moses, 1997) that found price changes during the introduction of new models to outnumber price changes for models that remain on the market. Stokey (1979) and Kahn (1986) argue that technologically
new models are often introduced at high price levels. After the introduction a rapid price fall occurs. In our study new items should be interpreted as new combinations of existing characteristics. With the exception of a DVD-player built into a television, a feature that we found occasionally at the end of the period, no really new technologies have been introduced. When price changes do not solely reflect differences in quality between new and disappearing models, this should be reflected by a systematic deviation of the residuals from the regression surface when prices are regressed on the characteristics of the different durables. If new models are based on new technologies, hidden price increases cannot always be detected by analysing residuals. The price change will presumably be caught by the new characteristic, if included in the regression.

For disappearing models a similar reasoning can be developed for relatively high as well for relatively low prices. New models i.e. new combinations of characteristics can make a durable good more effective, more powerful or less energy consuming. If this is not fully reflected in the price, 'older' models become relatively expensive. When these models are still on stock and the retailer wants to clear the shelves in favour of the new model he can choose to sell the 'old' model at sales' prices.

A detailed table with average residuals per period can be found in appendix 19. Table 14 contains some summary results. There is no evidence that all new and disappearing durable goods systematically are priced differently compared with similar goods that remain on the market. For three durable goods the prices of new items fall on average on the regression surface, only new televisions are slightly more expensive than similar already existing ones. A systematic difference may exist for disappearing televisions and PCs. During the period studied, prices of these items are structurally higher than their predicted values.

Table 14. Average residuals

|  | Televisions | Refrigerators | Washing <br> machines | Computers |
| :---: | :---: | :---: | :---: | :---: |
| Average residual of new models | 0.02 | 0.00 | 0.00 | 0.00 |
| Average residual of disappearing models | 0.05 | -0.01 | 0.00 | 0.03 |
| Average residual of new models (based on    <br> absolute values) 0.04 0.02 0.02 <br> Average residual of disappearing models    <br> (based on absolute values)    | 0.05 | 0.02 | 0.02 | 0.03 |

To get an impression of the average value of the residuals, whether positive or negative, the mean of the absolute values is also included in table 14 . For televisions this mean is highest, but does not exceed 0.05 . This points to a modest difference of approximately $5 \%$ between the observed prices and their predicted values.

Assumption ii in the previous section refers to the difference between the average residuals of new and disappearing items. Table 14 shows that these differences are quite
small. Table 15 compares $P_{G T(i)}$ with the second approximation, the matched-model Törnqvist index $P_{M T}$. Personal computers are the only durable good where an increasing spread is found between $P_{G T(i)}$ and $P_{M T}$, despite the fact that the difference between the average residuals of new and disappearing items is roughly the same for computers and televisions. What is important here, is the combination of the difference between the average residuals and the expenditure shares of matched items discussed in the previous section. Expression (7) shows that the matched-model expenditure share plays a decisive role in the effect of the difference between the average residuals. For refrigerators and washing machines this difference is very small and so is the difference between $P_{G T(i)}$ and $P_{M T}$.

Table 15. $P_{M T}$ and $P_{G T(i)}$, effect of assumption ii

| Period | Televisions |  | Refrigerators |  | Washing machines |  | Computers |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $P_{G T(i)}$ | $P_{M T}$ | $P_{G T(i)}$ | $P_{M T}$ | $P_{G T(i)}$ | $P_{M T}$ | $P_{G T(i)}$ | $P_{M T}$ |
| 199901 |  |  |  |  |  |  | 100.00 | 100.00 |
| 199902 | 100.00 | 100.00 | 100.00 | 100.00 | 100.00 | 100.00 | 96.02 | 94.77 |
| 199903 |  |  |  |  |  |  | 92.44 | 91.73 |
| 199904 | 97.01 | 96.95 | 99.02 | 99.09 | 98.01 | 97.96 | 88.66 | 89.17 |
| 199905 |  |  |  |  |  |  | 85.70 | 87.05 |
| 199906 | 94.73 | 94.87 | 97.81 | 98.06 | 97.87 | 97.56 | 82.42 | 84.05 |
| 199907 |  |  |  |  |  |  | 77.13 | 78.20 |
| 199908 | 92.99 | 93.25 | 97.51 | 97.68 | 98.18 | 97.93 | 70.68 | 72.51 |
| 199909 |  |  |  |  |  |  | 67.32 | 69.30 |
| 199910 | 90.75 | 91.10 | 97.49 | 97.26 | 97.73 | 97.54 | 67.64 | 69.69 |
| 199911 |  |  |  |  |  |  | 66.29 | 68.27 |
| 199912 | 89.44 | 89.79 | 96.86 | 96.79 | 97.15 | 96.94 | 64.95 | 67.61 |
| 200001 |  |  |  |  |  |  | 64.35 | 66.63 |
| 200002 | 88.24 | 88.72 | 96.23 | 96.26 | 96.50 | 96.19 | 63.24 | 66.37 |
| 200003 |  |  |  |  |  |  | 63.11 | 66.03 |
| 200004 | 87.61 | 87.63 | 96.02 | 95.61 | 96.00 | 95.66 | 62.18 | 65.33 |
| 200005 |  |  |  |  |  |  | 60.87 | 63.92 |
| 200006 | 87.07 | 87.31 | 95.15 | 94.64 | 95.18 | 94.79 | 57.21 | 60.55 |
| 200007 |  |  |  |  |  |  | 56.27 | 59.50 |
| 200008 | 86.34 | 86.63 | 94.24 | 93.77 | 93.86 | 93.64 | 55.09 | 58.59 |
| 200009 |  |  |  |  |  |  | 54.14 | 57.92 |
| 200010 | 85.27 | 85.77 | 94.02 | 93.72 | 93.15 | 93.00 | 52.65 | 56.47 |
| 200011 |  |  |  |  |  |  | 51.80 | 55.80 |
| 200012 | 85.01 | 85.42 | 94.22 | 93.61 | 92.69 | 92.47 | 51.48 | 55.01 |
| 200101 |  |  |  |  |  |  | 51.07 | 54.66 |
| 200102 | 85.34 | 85.71 | 95.09 | 94.27 | 92.93 | 92.62 | 48.80 | 52.61 |
| 200103 |  |  |  |  |  |  | 45.68 | 50.39 |
| 200104 | 85.48 | 85.70 | 95.76 | 95.09 | 93.64 | 93.44 | 45.38 | 50.02 |
| 200105 |  |  |  |  |  |  | 44.13 | 48.90 |
| 200106 | 85.52 | 85.76 | 95.97 | 94.46 | 93.46 | 93.03 | 42.77 | 47.69 |
| 200107 |  |  |  |  |  |  | 42.71 | 46.19 |
| 200108 | 84.56 | 84.77 | 94.8 | 93.57 | 92.73 | 92.57 | 40.37 | 44.51 |
| 200109 |  |  |  |  |  |  | 37.39 | 41.32 |
| 200110 | 83.68 | 84.10 | 94.07 | 93.41 | 92.16 | 92.06 | 36.86 | 40.50 |
| 200111 |  |  |  |  |  |  | 34.85 | 39.75 |
| 200112 | 82.97 | 83.44 | 94.85 | 93.20 | 91.39 | 91.21 | 33.95 | 39.02 |

Thus, for three of the four durable goods both assumptions approximately hold. This is again illustrated in figures 17-20, which depict the period-to-period changes of the generalised Törnqvist index and its approximations $P_{G T(i)}$ and $P_{M T}$.


Figure 17. Price indexes of televisions, changes with respect to the preceding period


Figure 18. Price indexes of refrigerators, changes with respect to the preceding period


Figure 19. Price indexes of washing machines, changes with respect to the preceding period


Figure 20. Price indexes of computers, changes with respect to the preceding period

### 4.4 Numerical comparison between the different indexes

The various price indexes can be compared by ranking them according to the distance between the index and the generalised Törnqvist index, since we consider the latter to be the reference or benchmark index. We define this distance as the average value of the absolute differences of the (bi)monthly indexes. The distances are given in table 16. For illustrative purposes we add the chained Laspeyres index. The first approximation of $P_{G T}$ is left out because this index will not be calculated in normal CPI practices.

Table 16. Distance between index and generalised Törnqvist index ${ }^{\text {a) }}$ b)

| Index | Televisions | Refrigerators | Washing <br> machines | Computers |
| :--- | :---: | :---: | :---: | :---: |
| $P_{M T}$ | $0.16(1)$ | $0.58(3)$ | $0.26(1)$ | $3.71(4)$ |
| Chained Laspeyres | $0.28(2)$ | $0.23(1)$ | $0.85(4)$ | $6.97(6)$ |
| OLS time dummy | $0.81(3)$ | $0.53(2)$ | $0.44(2)$ | $2.27(3)$ |
| WLS time dummy (quantities) | $2.18(5)$ | $0.75(4)$ | $0.81(3)$ | $0.72(1)$ |
| WLS time dummy (expenditure shares) | $1.67(4)$ | $2.22(5)$ | $1.70(5)$ | $1.18(2)$ |
| CPI | $2.73(6)$ | $2.87(6)$ | $2.58(6)$ | $5.64(5)$ |

${ }^{\text {a) }}$ Distance: average of absolute differences
${ }^{\text {b) }}$ Ranking of indexes is indicated between brackets
This table underlines the special position of computers. For televisions, refrigerators and washing machines the matched-model Törnqvist index performs best on average, although the chained Laspeyres is very close. Among the time dummy indexes the OLSversion approximates $P_{G T}$ clearly better than both weighted versions. Van der Grient and De Haan (2003) indicated that a time dummy price index should be completely based on observed prices when there are no new or disappearing goods. Timedependent weights, whether (relative) expenditures or quantities, do not satisfy that condition. This could be the underlying reason for the low ranking of both WLS time dummy indexes in table 16 . For computers, the ranking between matched-model and time dummy indexes is exactly the opposite. The quantity-weighted time dummy index performs best, while the chained Laspeyres is a bad candidate. The time dummy indexes for computers should be handled with some care, however. In section 4.1 we already expressed our doubts about the stability of the model parameters.

Generally spoken, the fraction of new and disappearing items seems to be decisive for a choice between matched-model and time dummy indexes. The distance between the CPI and the benchmark index is fairly big for all durables. The underlying causes were already discussed in section 4.2.

### 4.5 An aggregate hedonic index for 'durables'

The expenditures on the four durables discussed in this paper form nearly $50 \%$ of the CPI-weight of the total group of household appliances, audio-visual equipment and data processing equipment. An aggregate index of these four goods can thus considered to be
a reasonable approximation of an index of all durables (excluding cars). Based on the various price indexes presented above, we will calculate aggregate indexes. The weights are taken from the $\operatorname{CPI}(1995=100)$. Before, the monthly computer indexes have been recalculated on a bimonthly basis as unweighted averages of two monthly indexes. It is tempting to compare our aggregate indexes with the official CPI for all durables. However, this is complicated by the fact that the official CPI includes the (price reducing) effect of so-called energy-premiums for low-energy durables. Unfortunately, it proved impossible to calculate a CPI for all durables excluding those premiums. Therefore, we will restrict ourselves to a comparison of our aggregates, presented in table 17, with the CPI based on the four durables only. Figure 21 visualises the development of all aggregate indexes.

Table 17. Price index numbers for durables ${ }^{\text {a }}$ ( $199902=100$ )

|  | $P_{G T}$ | $P_{G T(i)}$ | $P_{M T}$ | WLS time <br> dummy ${ }^{\text {b }}$ | OLS time <br> dummy | CPI |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 199902 | 100.00 | 100.00 | 100.00 | 100.00 | 100.00 | 100.00 |
| 199904 | 95.66 | 95.52 | 95.71 | 95.46 | 95.36 | 96.73 |
| 199906 | 91.99 | 91.95 | 92.82 | 91.84 | 91.19 | 93.10 |
| 199908 | 87.28 | 87.20 | 88.07 | 87.89 | 88.15 | 88.48 |
| 199910 | 83.83 | 83.76 | 84.83 | 94.48 | 85.25 | 85.29 |
| 199912 | 82.40 | 82.41 | 83.62 | 83.38 | 83.56 | 82.57 |
| 200002 | 81.03 | 81.10 | 82.50 | 81.83 | 82.36 | 81.18 |
| 200004 | 80.14 | 80.32 | 81.66 | 80.85 | 81.13 | 79.21 |
| 200006 | 78.17 | 78.41 | 79.86 | 78.73 | 79.40 | 77.79 |
| 200008 | 76.20 | 76.46 | 78.02 | 77.14 | 76.97 | 75.45 |
| 200010 | 74.61 | 75.04 | 76.87 | 75.85 | 75.80 | 72.41 |
| 200012 | 73.78 | 74.17 | 75.90 | 75.04 | 75.29 | 71.84 |
| 200102 | 73.40 | 73.70 | 75.39 | 74.54 | 74.66 | 71.09 |
| 200104 | 71.78 | 72.11 | 74.15 | 72.98 | 73.27 | 70.97 |
| 200106 | 70.97 | 71.25 | 73.22 | 72.02 | 72.33 | 69.80 |
| 200108 | 69.39 | 69.90 | 71.50 | 70.63 | 70.64 | 68.58 |
| 200110 | 67.15 | 67.68 | 69.33 | 68.14 | 69.20 | 66.91 |
| 200112 | 65.59 | 66.22 | 68.32 | 66.47 | 68.02 | 66.31 |

${ }^{\text {a) }}$ Televisions, refrigerators, washing machines and computers ${ }^{\circ}$
${ }^{\text {b) }}$ Time-specific expenditure shares as weights.


Figure 21. Price indexes for durables $(199902=100)$

The difference between $P_{G T}$ and (its second approximation) $P_{M T}$ accumulates over time, up to slightly more than 3 indexpoints, which is nearly entirely due to computers. Apart from the period end 2000, early 2001, the CPI is surprisingly close to $P_{G T}$. The positive 'biases' for televisions, refrigerators and washing machines are more than compensated by the downward 'bias' for computers. However, it cannot be guaranteed that positive and negative 'biases' will compensate each other structurally.

## 5. Conclusion

This study had two aims: describing the market dynamics for some major durable goods and, most importantly, to investigate the potential of hedonic regression techniques for quality adjustments in the CPI. To this end we utilized scanner data on personal computers, televisions, washing machines and refrigerators. It appeared that all these durable goods have a clear seasonal pattern in their sales. Refrigerators reach their maximum sales level around August, the other three durables exhibit a peak at the end of the year. The market of each of these durables is dominated by approximately 10 brands. The distribution of sales is highly skewed, relatively few models count for the major part of total sales. The market is highly dynamic; each month many models appear on the market or disappear from it. Based on expenditures however, the attrition rate is much lower, except for computers.

The main focus in this paper is on hedonic modelling. Prices of televisions, refrigerators and washing machines can well be explained by the available technical characteristics and dummy variables for brand name and outlet type. Using OLS, the coefficients for the most important characteristics appeared to be stable for at least half a year. This makes hedonics a promising tool for quality adjustments in the CPI, the advantage being that the data collection on characteristics is needed only twice a year. The lack of weights is less important when an explicit variant of the hedonic technique will be used.

For implicit quality adjustments, like the time dummy variant, weighting is essential. The data requirements, however, are stronger because prices, quantities and of course characteristics are needed on a monthly basis. Using WLS, the coefficients appear to be a bit less stable. Parameter stability was satisfactory between adjacent periods, though. The use of different sets of weights resulted in substantial differences in the resulting indexes. We took the WLS-procedure proposed by Diewert (2003), leading to a generalised Törnqvist index, as our benchmark. The OLS time dummy index did not differ much from the generalised Törnqvist, which was actually surprising. We also compared the official CPI-figures with this benchmark price index. For televisions, washing machines and refrigerators the CPI was systematically above the generalised Törnqvist index, for computers the situation was reverse. Sampling aspects seem to be a major reason in explaining these differences.

Finally, we investigated whether a superlative matched-model index would suffice or new and disappearing items should explicitly be taken into account. Our findings confirmed that this is mainly determined by market dynamics. As the attrition rates, based on expenditures, are low for televisions, washing machines and refrigerators the matched-model Törnqvist index appeared to be a very good approximation of the generalised one. For computers, where the expenditure shares of new and disappearing items are substantial, a matched-model index seems not to suffice.

Our conclusion that a matched-model index suffices in case the attrition rates of new and disappearing items are low, is not very surprising. Less evident was our finding that there seemed to be no structural pricing policy for new and disappearing goods. An exception could be observed for disappearing televisions and PCs. The price level of these items seems to be structurally higher than the price level of similar models that remain on the market.

Although the use of hedonics seems promising, the practical application is not a simple next step. One problem is the use of brand names in the regression model. Since brand effects are important it could be advisable (data permitting) to estimate hedonic models for individual brands or for groups of similar brands. Secondly, rules for how to deal with outliers should be developed, a point that has earlier been stressed by Silver (2002). Thirdly, the bias possibly created by the seasonal pattern in sales of durables should be examined. Last but not least, sampling remains an important issue. For the time being it is not realistic to expect that on a regular basis population data, like the scanner data we utilized, can be obtained. Sampling aspects should receive more attention in research on hedonics since this is where improvements are feasible.

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## Appendix 1. Representativity of CPI-sample

As stated in section 2 of the main text, the scanner data can be viewed as the population of sales over the period 1999-2001. This enabled us to examine whether the samples used for the official CPI were representative for the population. Table 1-1 indicates whether Statistics Netherlands has been successful in selecting the best selling models for their price collection. All models available during 1999-2001 are ranked according to their numbers sold. We distinguish four groups, the top 10 of models, two groups with intermediate sales and a last group of models with modest sales. Table 1-1 shows how these groups are represented in the CPI-sample.

Table 1-1. Representativity of CPI-sample, 1999-2001 ${ }^{\text {a) }}$

| Ranking of models in population, | Number of models in CPI-sample |  |  |
| :---: | :---: | :---: | :---: |
| based on numbers sold | Televisions | Refrigerators | Washing machines |
| $1-10$ | 6 | 0 | 0 |
| $11-25$ | 2 | 2 | 2 |
| $26-100$ | 7 | 6 | 2 |
| $>=101$ | 20 | 23 | 18 |
| Total | 35 | 31 | 22 |
| a) Computers are not included, due to the specific market with many 'clones’ |  |  |  |

The information should be handled with some care. The grouping is made at the level of individual models and not on brands. For televisions for instance, the top 25 of best sold models is dominated by two brands. These brands are well represented. To represent other brands as well, automatically models are selected which are lower on the list.

The table shows that Statistics Netherlands has not always been successful in tracking the dynamics of the market. For both refrigerators and washing machines the ten best sold models were not included in the CPI-sample. For refrigerators this can partly be explained by the structure of the market. The list of best-sold refrigerators is dominated by a few models of a less important brand. The brands which dominate the market all sell a wide variety of models, each with medium sales. As Statistics Netherlands' main concern is to have the most important brands to be represented in their sample, it is clear why the best sold models are not selected. The brands that dominate the market are all included in the CPI-sample although not always with their best sold models. One reason could be that updating of the sample is not considered urgent as long as enough price observations (for older but still existing models) are available. A second reason is that the replacement rule for an item that has become obsolete - i.e. choosing the most comparable item - might result in selecting the next obsolete item. However, this is common practice in many statistical offices. Silver and Heravi (2002) and Schultze and Mackie (2001) argue that sample degradation can be a source of bias. The criterion for selecting a replacement item should be representativity rather than comparability.

## Appendix 2. Televisions: List of all characteristics used

| Variable | Explanation | Base model |
| :--- | :--- | :--- |
| Brand01 to brand77 | Brand | Brand05 |
| Outlet2 | Sold in chain stores | Sold in buying combinations |
| Outlet3 | Sold in department stores or mail-order houses | Sold in buying combinations |
| Outlet4 | Sold in independent stores | Sold in buying combinations |
| Outlet5 | Sold in photo retail stores | Sold in buying combinations |
| Diameter | Screen diameter |  |
| Scrtype2 | Flat screen | Spherical screen |
| Ntuners | Television with 2 tuners | Television with 1 tuner |
| Pip1 | Option 'picture in picture' present | No 'picture in picture’ |
| Mltpip1 | Dynamic picture and text | No multipip possibilities |
| Mltpip2 | Dynamic picture and picture | No multipip possibilities |
| Syst2 | Pal / Secam | Pal |
| Syst4 | Pal / Secam / ntsc | Pal |
| Syst6 | Pal + | Pal |
| Syst7 | Pal + Secam | Pal |
| Syst8 | Pal + / Secam / ntsc | Pal |
| Model | Speakers on side | Frontal speakers |
| Freeze1 | Possibility to freeze picture | No possibility to freeze picture |
| S_vhs1 | s-vhs connector present (for camcorder) | No s-vhs connector |
| Sattune1 | Built-in satellite tuner | No satellite tuner |
| Nicam1 | Nicam decoder present | No Nicam decoder |
| Scratio2 | Widescreen (16:9) | Standard screen (4:3) |
| Freq100 | 100 Hz | 50 Hz |
| Wattage | Power |  |
| Dolby1 | Dolby surround sound | No Dolby surround sound |
| Dvd | Presence of DVD-player | No DVD-player |
| Portab | Portable Television | Not portable |
| Remote0 | No remote control | With remote control |
| Sound0 | Mono sound | Stereo sound |
| Text0 | No teletext | Teletext with toptext |
| Text1 | Simple teletext | Teletext with toptext |
| Text3 | Teletext with 2.5 resolution | Teletext with toptext |
| Pc_conn | PC connection | No PC connection |
| Flatscr | Completely flat screen | No completely flat screen |
|  |  |  |

## Appendix 3. Refrigerators: List of all characteristics used

| Variable | Explanation | Base model |
| :--- | :--- | :--- |
| Brand01 to brand76 | Brand | Brand05 |
| Outlet2 | Sold in chain stores | Sold in buying combinations |
| Outlet3 | Sold in department stores or mail-order houses | Sold in buying combinations |
| Outlet4 | Sold in independent stores | Sold in buying combinations |
| Outlet5 | Sold in kitchen retail stores | Sold in buying combinations |
| Doors_1,_3 and_4 | Number of doors | Two doors |
| No_frost | No-frost option present | No-frost option not present |
| Combi | Fridge freezer | Refrigerator |
| Star_1,_2 and_3 | Refrigerator with 1,2 or 3 stars | Refrigerator with 4 stars |
| Groscool | Gross content cooling part (litres) |  |
| Grosfrez | Gross content freezing part (litres) |  |
| No-kwh | Energy-use |  |
| Syst_0 | Absorption system |  |
| Syst_2 | Two compressors | One compressor |
| No_label | No low-energy refrigerator | One compressor |
| X_label | Low-energy refrigerator, class not specified | Low-energy refrigerator, Class A |
| B_label | Low-energy refrigerator, class B | Low-energy refrigerator, Class A |
| C_label | Low-energy refrigerator, class C | Low-energy refrigerator, Class A |
| D_label | Low-energy refrigerator, class D | Low-energy refrigerator, Class A |
| E_label | Low-energy refrigerator, class E | Low-energy refrigerator, Class A |
| F_label | Low-energy refrigerator, class F | Low-energy refrigerator, Class A |
| G_label | Low-energy refrigerator, class G | Low-energy refrigerator, Class A |
| A1_label | Low-energy refrigerator, class A-plus | Low-energy refrigerator, Class A |

## Appendix 4. Washing machines: List of all characteristics used

| Variable | Explanation | Base model |
| :--- | :--- | :--- |
| Rotspeed | Rotation speed |  |
| Type_cw | Separate built-in spin-dryer | No separate spin-dryer |
| Type_1 | Top loader | Frontloaded |
| Brand01 to brand45 | Brand | Brand 01 |
| EnerA to EnerF | Energy label A to F | Energy label A |
| Centr_m | Mechanical spin-dryer | Electronical spin-dryer |
| WashresA to WashresG | Washing result A to G | Washing result A |
| Buying | Sold in buying combinations | Sold in chain stores |
| Depmoh | Sold in department stores or mail-order houses | Sold in chain stores |
| Indep | Sold in independent stores | Sold in chain stores |

## Appendix 5. Personal computers: List of all characteristics used

| Variable | Explanation | Base model |
| :--- | :--- | :--- |
| Speed | Processor speed (MHz) |  |
| Hdisk | Storage capacity of hard disk (MB) |  |
| Memory | Memory capacity (MB) | No monitor |
| Monitor | Presence of monitor | Stand alone PC |
| Workstat | Stand alone PC or workstation in Network | No USB port |
| USB | USB port present |  |
| Proc01 to proc27 | Type of processor | Sold in computer stores |
| Brand01 to brand34 | Brand | Sold in computer stores |
| Buying | Sold in buying combinations | Sold in computer stores |
| Chains | Sold in chain stores | Sold in computer stores |
| Depmoh | Sold in department stores or mail-order houses | Sold in computer stores |
| Indep | Sold in independent stores | Sold in computer stores |
| Prt | Sold in Photo retail stores |  |
| Tcs | Sold in Telecom specialists shops |  |


| 免 |  | $B_{3}^{\infty}$ |  |  |  |  |  |  |  |  |  |  | yol | in | 感 |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 衣 | \％ | \％ | 为象 | － | 通 |  | ， | ${ }^{\circ}$ |  | 年 | \％ |  |  | $b_{0}^{0}$ |  |  |  | \％ |
| $\left\lvert\, \frac{0}{\frac{0}{5}}\right.$ |  | blo |  | 答 | 通 |  |  |  | 능 \％\％ |  | \％ | ${ }^{\circ}$ | \％ | $\square^{\circ}$ |  |  |  |  |
|  |  | on on ex ex |  | O | ${ }^{\circ}$ |  |  |  | $\circ$ | O0\％ | ${ }_{0}^{\circ}$ | － | 部栜 | ${ }^{\circ}$ |  |  |  |  |
|  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
|  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  | $\begin{array}{\|l\|} \hline \end{array}$ |  | 范 |




1）Coefficients in bold are significant at the $5 \%$ level in all periods
Appendix 8. Washing machines: Summary results of 18 bimonthly OLS-regressions 1)

|  |  | Average | st dev | min | max |
| :---: | :---: | :---: | :---: | :---: | :---: |
| brand23 | coefficient st. error | $\begin{array}{r} -0,379 \\ 0,101 \\ \hline \end{array}$ | $\begin{aligned} & \hline 0,064 \\ & 0,047 \\ & \hline \end{aligned}$ | $\begin{aligned} & -0,465 \\ & 0,049 \\ & \hline \end{aligned}$ | $\begin{aligned} & \hline-0,288 \\ & 0,207 \\ & \hline \end{aligned}$ |
| brand24 | coefficient st. error | $\begin{aligned} & -0,185 \\ & 0,111 \\ & \hline \end{aligned}$ | $\begin{aligned} & \hline 0,057 \\ & 0,024 \\ & \hline \end{aligned}$ | $\begin{aligned} & -0,274 \\ & 0,071 \\ & \hline \end{aligned}$ | $\begin{aligned} & -0,127 \\ & 0,130 \\ & \hline \end{aligned}$ |
| brand25 | coefficient <br> st. error | $\begin{aligned} & -0,337 \\ & 0,048 \\ & \hline \end{aligned}$ | $\begin{aligned} & \hline 0,068 \\ & 0,009 \\ & \hline \end{aligned}$ | $\begin{aligned} & \hline-0,458 \\ & 0,030 \\ & \hline \end{aligned}$ | $\begin{gathered} \hline-0,199 \\ 0,066 \\ \hline \end{gathered}$ |
| brand26 | coefficient st. error | $\begin{aligned} & -0,456 \\ & 0,120 \\ & \hline \end{aligned}$ | $\begin{array}{r} \hline 0,097 \\ 0,018 \\ \hline \end{array}$ | $\begin{aligned} & \hline-0,606 \\ & 0,087 \\ & \hline \end{aligned}$ | $\begin{aligned} & \hline-0,351 \\ & 0,143 \\ & \hline \end{aligned}$ |
| brand27 | coefficient st. error | $\begin{aligned} & -0,388 \\ & 0,074 \\ & \hline \end{aligned}$ | $\begin{aligned} & 0,056 \\ & 0,023 \\ & \hline \end{aligned}$ | $\begin{aligned} & -0,491 \\ & 0,051 \\ & \hline \end{aligned}$ | $\begin{aligned} & -0,296 \\ & 0,137 \\ & \hline \end{aligned}$ |
| brand28 | coefficient st. error | $\begin{aligned} & -\mathbf{- 0 , 4 7 5} \\ & 0,027 \\ & \hline \end{aligned}$ | $\begin{aligned} & \hline 0,029 \\ & 0,002 \\ & \hline \end{aligned}$ | $\begin{aligned} & \hline-0,540 \\ & 0,020 \\ & \hline \end{aligned}$ | $\begin{aligned} & \hline-0,414 \\ & 0,031 \\ & \hline \end{aligned}$ |
| brand29 | coefficient <br> st. error | $\begin{aligned} & -0,083 \\ & 0,067 \\ & \hline \end{aligned}$ | $\begin{aligned} & \hline 0,046 \\ & 0,031 \\ & \hline \end{aligned}$ | $\begin{aligned} & \hline-0,164 \\ & 0,043 \\ & \hline \end{aligned}$ | $\begin{aligned} & \hline 0,013 \\ & 0,142 \\ & \hline \end{aligned}$ |
| brand30 | coefficient st. error | $\begin{aligned} & \hline-\mathbf{0 , 1 2 4} \\ & 0,036 \\ & \hline \end{aligned}$ | $\begin{aligned} & 0,021 \\ & 0,005 \\ & \hline \end{aligned}$ | $\begin{aligned} & \hline-0,167 \\ & 0,030 \\ & \hline \end{aligned}$ | $\begin{aligned} & -0,092 \\ & 0,044 \\ & \hline \end{aligned}$ |
| brand31 | coefficient st. error | $\begin{aligned} & 0,297 \\ & 0,018 \\ & \hline \end{aligned}$ | $\begin{aligned} & \hline 0,026 \\ & 0,001 \\ & \hline \end{aligned}$ | $\begin{array}{r} \hline 0,258 \\ 0,014 \\ \hline \end{array}$ | $\begin{aligned} & \hline 0,347 \\ & 0,019 \\ & \hline \end{aligned}$ |
| brand32 | coefficient <br> st. error | $\begin{aligned} & -0,332 \\ & 0,038 \\ & \hline \end{aligned}$ | $\begin{aligned} & \hline 0,038 \\ & 0,003 \\ & \hline \end{aligned}$ | $\begin{aligned} & \hline-0,392 \\ & 0,030 \\ & \hline \end{aligned}$ | $\begin{aligned} & \hline-0,267 \\ & 0,047 \\ & \hline \end{aligned}$ |
| brand33 | coefficient st. error | $\begin{aligned} & -0,446 \\ & 0,108 \\ & \hline \end{aligned}$ | $\begin{aligned} & \hline 0,074 \\ & 0,018 \\ & \hline \end{aligned}$ | $\begin{aligned} & \hline-0,579 \\ & 0,083 \\ & \hline \end{aligned}$ | $\begin{aligned} & \hline-0,283 \\ & 0,141 \\ & \hline \end{aligned}$ |
| brand34 | coefficient st. error | $\begin{aligned} & \hline-0,718 \\ & 0,104 \\ & \hline \end{aligned}$ | $\begin{array}{r} \hline 0,076 \\ 0,024 \\ \hline \end{array}$ | $\begin{aligned} & -0,900 \\ & 0,070 \\ & \hline \end{aligned}$ | $\begin{aligned} & \hline-0,576 \\ & 0,141 \\ & \hline \end{aligned}$ |
| brand35 | coefficient <br> st. error | $\begin{aligned} & -0,495 \\ & 0,094 \\ & \hline \end{aligned}$ | $\begin{aligned} & \hline 0,066 \\ & 0,032 \\ & \hline \end{aligned}$ | $\begin{aligned} & \hline-0,609 \\ & 0,051 \\ & \hline \end{aligned}$ | $\begin{aligned} & \hline-0,386 \\ & 0,143 \\ & \hline \end{aligned}$ |
| brand36 | coefficient st. error | $\begin{aligned} & \hline-0,137 \\ & 0,113 \\ & \hline \end{aligned}$ | $\begin{array}{r} \hline 0,138 \\ 0,021 \\ \hline \end{array}$ | $\begin{aligned} & \hline-0,357 \\ & 0,091 \\ & \hline \end{aligned}$ | $\begin{aligned} & \hline 0,024 \\ & 0,145 \\ & \hline \end{aligned}$ |
| brand37 | coefficient <br> st. error | $\begin{aligned} & -0,344 \\ & 0,032 \\ & \hline \end{aligned}$ | $\begin{aligned} & \hline 0,041 \\ & 0,003 \\ & \hline \end{aligned}$ | $\begin{aligned} & \hline-0,406 \\ & 0,028 \\ & \hline \end{aligned}$ | $\begin{aligned} & \hline-0,243 \\ & 0,038 \\ & \hline \end{aligned}$ |
| brand38 | coefficient st. error | $\begin{aligned} & \hline-0,006 \\ & 0,019 \\ & \hline \end{aligned}$ | $\begin{aligned} & \hline 0,038 \\ & 0,001 \\ & \hline \end{aligned}$ | $\begin{aligned} & \hline-0,058 \\ & 0,015 \\ & \hline \end{aligned}$ | $\begin{aligned} & \hline 0,074 \\ & 0,021 \\ & \hline \end{aligned}$ |
| brand39 | coefficient st. error | $\begin{aligned} & \hline-0,147 \\ & 0,086 \\ & \hline \end{aligned}$ | $\begin{aligned} & \hline 0,230 \\ & 0,011 \\ & \hline \end{aligned}$ | $\begin{aligned} & \hline-0,543 \\ & 0,072 \\ & \hline \end{aligned}$ | $\begin{aligned} & \hline 0,014 \\ & 0,101 \\ & \hline \end{aligned}$ |
| brand41 | coefficient st. error | $\begin{aligned} & -0,151 \\ & 0,123 \\ & \hline \end{aligned}$ | $\begin{aligned} & \hline 0,142 \\ & 0,028 \\ & \hline \end{aligned}$ | $\begin{aligned} & -0,362 \\ & 0,082 \\ & \hline \end{aligned}$ | $\begin{aligned} & \hline 0,274 \\ & 0,191 \\ & \hline \end{aligned}$ |
| brand42 | coefficient st. error | $\begin{aligned} & -0,257 \\ & 0,023 \\ & \hline \end{aligned}$ | $\begin{aligned} & \hline 0,023 \\ & 0,002 \\ & \hline \end{aligned}$ | $\begin{aligned} & -0,294 \\ & 0,020 \\ & \hline \end{aligned}$ | $\begin{aligned} & -0,211 \\ & 0,026 \\ & \hline \end{aligned}$ |
| brand43 | coefficient st. error | $\begin{aligned} & -0,075 \\ & 0,031 \\ & \hline \end{aligned}$ | $\begin{aligned} & \hline 0,030 \\ & 0,004 \\ & \hline \end{aligned}$ | $\begin{aligned} & \hline-0,142 \\ & 0,022 \\ & \hline \end{aligned}$ | $\begin{aligned} & \hline-0,021 \\ & 0,036 \\ & \hline \end{aligned}$ |
| brand44 | coefficient st. error | $\begin{aligned} & -0,234 \\ & 0,023 \\ & \hline \end{aligned}$ | $\begin{aligned} & \hline 0,029 \\ & 0,002 \\ & \hline \end{aligned}$ | $\begin{aligned} & \hline-0,283 \\ & 0,017 \\ & \hline \end{aligned}$ | $\begin{aligned} & \hline-0,175 \\ & 0,026 \\ & \hline \end{aligned}$ |
| brand45 | coefficient st. error | $\begin{aligned} & -0,463 \\ & 0,141 \\ & \hline \end{aligned}$ | $\begin{aligned} & 0,000 \\ & 0,000 \\ & \hline \end{aligned}$ | $\begin{aligned} & -0,463 \\ & 0,141 \\ & \hline \end{aligned}$ | $\begin{aligned} & \hline-0,463 \\ & 0,141 \\ & \hline \end{aligned}$ |


|  |  | Average | st dev | min | max |
| :---: | :---: | :---: | :---: | :---: | :---: |
| brand02 | coefficient st. error | $\begin{aligned} & \hline-0,643 \\ & 0,111 \\ & \hline \end{aligned}$ | $\begin{array}{r} \hline 0,088 \\ 0,021 \\ \hline \end{array}$ | $\begin{aligned} & \hline-0,754 \\ & 0,087 \\ & \hline \end{aligned}$ | $\begin{aligned} & -0,494 \\ & 0,145 \\ & \hline \end{aligned}$ |
| brand03 | coefficient <br> st. error | $\begin{aligned} & \hline-0,320 \\ & 0,046 \\ & \hline \end{aligned}$ | $\begin{aligned} & \hline 0,055 \\ & 0,008 \\ & \hline \end{aligned}$ | $\begin{aligned} & \hline-0,402 \\ & 0,027 \\ & \hline \end{aligned}$ | $\begin{aligned} & \hline-0,207 \\ & 0,060 \\ & \hline \end{aligned}$ |
| brand04 | coefficient st. error | $\begin{array}{r} \mathbf{0 , 2 8 0} \\ 0,050 \\ \hline \end{array}$ | $\begin{array}{r} \hline 0,044 \\ 0,016 \\ \hline \end{array}$ | $\begin{aligned} & \hline 0,221 \\ & 0,036 \\ & \hline \end{aligned}$ | $\begin{array}{r} \hline 0,373 \\ 0,096 \\ \hline \end{array}$ |
| brand05 | coefficient st. error | $\begin{aligned} & -0,887 \\ & 0,136 \\ & \hline \end{aligned}$ | $\begin{array}{r} 0,000 \\ 0,000 \\ \hline \end{array}$ | $\begin{aligned} & -0,887 \\ & 0,136 \\ & \hline \end{aligned}$ | $\begin{array}{r} -0,887 \\ 0,136 \\ \hline \end{array}$ |
| brand06 | coefficient <br> st. error | $\begin{aligned} & \hline-0,225 \\ & 0,078 \\ & \hline \end{aligned}$ | $\begin{aligned} & \hline 0,057 \\ & 0,004 \\ & \hline \end{aligned}$ | $\begin{aligned} & \hline-0,331 \\ & 0,070 \\ & \hline \end{aligned}$ | $\begin{array}{r} \hline-0,134 \\ 0,085 \\ \hline \end{array}$ |
| brand07 | coefficient <br> st. error | $\begin{aligned} & \hline-0,105 \\ & 0,019 \\ & \hline \end{aligned}$ | $\begin{aligned} & \hline 0,016 \\ & 0,001 \\ & \hline \end{aligned}$ | $\begin{aligned} & \hline-0,143 \\ & 0,016 \\ & \hline \end{aligned}$ | $\begin{aligned} & \hline-0,078 \\ & 0,022 \\ & \hline \end{aligned}$ |
| brand08 | coefficient st. error | $\begin{aligned} & -0,410 \\ & 0,085 \\ & \hline \end{aligned}$ | $\begin{array}{r} 0,049 \\ 0,030 \\ \hline \end{array}$ | $\begin{aligned} & \hline-0,497 \\ & 0,059 \\ & \hline \end{aligned}$ | $\begin{aligned} & -0,309 \\ & 0,142 \\ & \hline \end{aligned}$ |
| brand09 | coefficient <br> st. error | $\begin{aligned} & \hline 0,069 \\ & 0,082 \\ & \hline \end{aligned}$ | $\begin{aligned} & \hline 0,121 \\ & 0,037 \\ & \hline \end{aligned}$ | $\begin{aligned} & \hline-0,284 \\ & 0,031 \\ & \hline \end{aligned}$ | $\begin{aligned} & \hline 0,194 \\ & 0,142 \\ & \hline \end{aligned}$ |
| brand10 | coefficient <br> st. error | $\begin{aligned} & -0,616 \\ & 0,119 \\ & \hline \end{aligned}$ | $\begin{aligned} & \hline 0,138 \\ & 0,020 \\ & \hline \end{aligned}$ | $\begin{aligned} & \hline-0,830 \\ & 0,088 \\ & \hline \end{aligned}$ | $\begin{aligned} & -0,411 \\ & 0,145 \\ & \hline \end{aligned}$ |
| brand11 | coefficient st. error | $\begin{aligned} & -0,012 \\ & 0,018 \\ & \hline \end{aligned}$ | $\begin{aligned} & \hline 0,027 \\ & 0,001 \\ & \hline \end{aligned}$ | $\begin{gathered} -0,045 \\ 0,014 \\ \hline \end{gathered}$ | $\begin{aligned} & \hline 0,041 \\ & 0,019 \\ & \hline \end{aligned}$ |
| brand12 | coefficient <br> st. error | $\begin{aligned} & -0,083 \\ & 0,035 \\ & \hline \end{aligned}$ | $\begin{aligned} & \hline 0,048 \\ & 0,011 \\ & \hline \end{aligned}$ | $\begin{aligned} & \hline-0,164 \\ & 0,025 \\ & \hline \end{aligned}$ | $\begin{aligned} & \hline 0,009 \\ & 0,070 \\ & \hline \end{aligned}$ |
| brand13 | coefficient <br> st. error | $\begin{aligned} & -0,317 \\ & 0,028 \\ & \hline \end{aligned}$ | $\begin{aligned} & \hline 0,026 \\ & 0,002 \\ & \hline \end{aligned}$ | $\begin{gathered} \hline-0,370 \\ 0,024 \\ \hline \end{gathered}$ | $\begin{aligned} & \hline-0,279 \\ & 0,031 \\ & \hline \end{aligned}$ |
| brand14 | coefficient st. error | $\begin{aligned} & -0,617 \\ & 0,125 \\ & \hline \end{aligned}$ | $\begin{aligned} & \hline 0,053 \\ & 0,021 \\ & \hline \end{aligned}$ | $\begin{gathered} \hline-0,696 \\ 0,092 \\ \hline \end{gathered}$ | $\begin{aligned} & -0,534 \\ & 0,147 \\ & \hline \end{aligned}$ |
| brand15 | coefficient <br> st. error | $\begin{aligned} & -0,305 \\ & 0,108 \\ & \hline \end{aligned}$ | $\begin{aligned} & \hline 0,098 \\ & 0,024 \\ & \hline \end{aligned}$ | $\begin{aligned} & \hline-0,396 \\ & 0,070 \\ & \hline \end{aligned}$ | $\begin{aligned} & \hline-0,126 \\ & 0,145 \\ & \hline \end{aligned}$ |
| brand16 | coefficient <br> st. error | $\begin{aligned} & -0,340 \\ & 0,099 \\ & \hline \end{aligned}$ | $\begin{aligned} & \hline 0,014 \\ & 0,002 \\ & \hline \end{aligned}$ | $\begin{aligned} & \hline-0,364 \\ & 0,096 \\ & \hline \end{aligned}$ | $\begin{aligned} & -0,327 \\ & 0,100 \\ & \hline \end{aligned}$ |
| brand17 | coefficient st. error | $\begin{aligned} & -0,366 \\ & 0,038 \\ & \hline \end{aligned}$ | $\begin{aligned} & \hline 0,040 \\ & 0,003 \\ & \hline \end{aligned}$ | $\begin{gathered} -0,404 \\ 0,036 \\ \hline \end{gathered}$ | $\begin{aligned} & -0,293 \\ & 0,045 \\ & \hline \end{aligned}$ |
| brand18 | coefficient <br> st. error | $\begin{aligned} & -0,419 \\ & 0,073 \\ & \hline \end{aligned}$ | $\begin{aligned} & \hline 0,109 \\ & 0,027 \\ & \hline \end{aligned}$ | $\begin{aligned} & \hline-0,594 \\ & 0,046 \\ & \hline \end{aligned}$ | $\begin{aligned} & \hline-0,199 \\ & 0,141 \\ & \hline \end{aligned}$ |
| brand19 | coefficient <br> st. error | $\begin{aligned} & \hline-0,384 \\ & 0,028 \\ & \hline \end{aligned}$ | $\begin{aligned} & \hline 0,047 \\ & 0,002 \\ & \hline \end{aligned}$ | $\begin{aligned} & \hline-0,482 \\ & 0,021 \\ & \hline \end{aligned}$ | $\begin{aligned} & \hline-0,319 \\ & 0,031 \\ & \hline \end{aligned}$ |
| brand20 | coefficient <br> st. error | $\begin{aligned} & -0,216 \\ & 0,081 \\ & \hline \end{aligned}$ | $\begin{aligned} & \hline 0,104 \\ & 0,028 \\ & \hline \end{aligned}$ | $\begin{aligned} & \hline-0,359 \\ & 0,051 \\ & \hline \end{aligned}$ | $\begin{aligned} & \hline-0,047 \\ & 0,136 \\ & \hline \end{aligned}$ |
| brand21 | coefficient <br> st. error | $\begin{aligned} & \hline 0,384 \\ & 0,138 \\ & \hline \end{aligned}$ | $\begin{aligned} & \hline 0,023 \\ & 0,002 \\ & \hline \end{aligned}$ | $\begin{aligned} & \hline 0,361 \\ & 0,135 \\ & \hline \end{aligned}$ | $\begin{aligned} & \hline 0,406 \\ & 0,140 \\ & \hline \end{aligned}$ |
| brand22 | coefficient <br> st. error | $\begin{aligned} & \hline-0,288 \\ & 0,073 \\ & \hline \end{aligned}$ | $\begin{aligned} & \hline 0,092 \\ & 0,010 \\ & \hline \end{aligned}$ | $\begin{aligned} & \hline-0,420 \\ & 0,059 \\ & \hline \end{aligned}$ | $\begin{aligned} & \hline-0,169 \\ & 0,096 \\ & \hline \end{aligned}$ |


| ๔ | $\left\|\begin{array}{cc} \mathscr{\infty} & 0 \\ 20 & \vdots \\ i n & 0 \end{array}\right\|$ | $\begin{array}{\|cc\|} \hline 8 & 0 \\ 0-8 \\ 0 \end{array}$ |  | $\left.\begin{array}{\|cc\|} \hline 0 & n \\ \frac{n}{0} & \stackrel{0}{0} \end{array}\right\}$ | $0$ |  | $\begin{array}{ll} \hline & 0 \\ \hline & 0 \\ \hline \end{array}$ |  | $\left\|\begin{array}{cc} \begin{array}{c} 7 \\ \hline \end{array} & 0 \\ \hdashline & 0 \\ 0 \end{array}\right\|$ |  |  | $0$ |  |  | $\frac{N}{2}$ |  |  |  |  | $1 \begin{array}{lll} \hline 8 & 0 \\ \hline & \vdots \\ 0 & 0 \end{array}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| ᄃ | $\left.\begin{array}{\|cc\|} \hline 9 & 0 \\ 6 & 0 \\ 0 & 0 \\ 10 & 0 \end{array} \right\rvert\,$ | $\begin{array}{\|cc\|} \hline-8 & 0 \\ \hline 0.0 \\ \hline 0 & 0 \end{array}$ |  | $\begin{array}{\|cc\|} \hline 0.0 \\ \hline 0 & \overline{0} \\ \hline \end{array}$ | $\left.\begin{array}{\|cc\|} \hline & 0 \\ 8 & 0 \\ 0 & 0 \end{array} \right\rvert\,$ | $\begin{array}{\|cc\|} \hline 0 & 0 \\ \hline 0 & 0 \\ 0 & 0 \end{array}$ | $\left\|\begin{array}{cc} 1 & - \\ \hline & y \\ 0 & 0 \end{array}\right\|$ |  | $\begin{array}{ll} \hline 8 & 7 \\ 0 & 7 \\ 0 & 0 \end{array}$ |  | $5$ |  |  | $\left\lvert\, \begin{array}{cc} \hline 8 \\ \hline \end{array}\right.$ |  | $\begin{array}{ll} \hline 0 & 0 \\ 0 & 0 \\ 0 & 0 \\ i \end{array}$ |  | $\begin{array}{\|cc\|} \hline 0 & 0 \\ \hline-0 & 0 \\ \hline-0 & 0 \end{array}$ | $\begin{array}{lll} \hline 0 & 0 \\ \hline 0 & 0 \\ 0 & 0 \end{array}$ | $\begin{array}{ll} 1 & 0 \\ 0 & 0 \\ 0 & 0 \end{array}$ |
| O | $\left.\begin{array}{\|cc\|} \hline N & N \\ 0 & 0 \\ 0 & 0 \end{array} \right\rvert\,$ | $0$ | $5$ | $\left.\begin{array}{\|cc\|} \hline 0 & 0 \\ 0 & 0 \\ 0 & 0 \\ 0 \end{array} \right\rvert\,$ | $0$ | $\begin{array}{\|cc\|} \hline 20 & 0 \\ \hline & 0 \\ 0 & 0 \end{array}$ | $\left\|\begin{array}{cc} -2 & 0 \\ 0 & 0 \\ 0 & 0 \end{array}\right\|$ | $\mathfrak{l l}$ | $\left\|\begin{array}{ll} \hat{0} & \hat{0} \\ 0 & 0 \\ 0 & 0 \end{array}\right\|$ | $\left.\right\|_{0} ^{0}$ |  | $\begin{array}{lll} \hline 10 & 0 \\ 0 & 0 \\ 0 & 0 \\ -1 & 0 \end{array}$ |  |  | $\frac{0}{2}$ | $0_{5}^{N}$ |  | $\begin{array}{\|cc\|} \hline \begin{array}{c} 2 \\ \hline \end{array} & \overline{0} \\ 0 & 0 \end{array}$ | $\left\lvert\, \begin{array}{\|cc\|} \hline 0 & 2 \\ 0 & 0 \\ 0 & 0 \\ \hline \end{array}\right.$ | $\begin{array}{\|cc\|} \hline 10 & 2 \\ \hline & 0 \\ 0 & 0 \\ \hline \end{array}$ |
| \% |  | $0$ |  | $\left.\begin{array}{\|cc\|} \hline 0 & 0 \\ \hline & \underset{0}{0} \\ \hline \end{array} \right\rvert\,$ | $\left\lvert\, \begin{array}{ll} \hline 0 & N \\ 0 & N \\ 0 & 0 \end{array}\right.$ |  | $\left\|\begin{array}{ll} \hline 0 \\ \hline 0 & 0 \\ 0 & 0 \\ 0 & 0 \end{array}\right\|$ | $3 \begin{array}{ll} \infty & 0 \\ n & 0 \\ 0 & 0 \\ 0 & 0 \end{array}$ | $\left\|\begin{array}{\|cc\|} \hline \begin{array}{l} 0 \\ \hline \end{array} \\ \hline & \mathbf{m} \\ \hline \end{array}\right\|$ |  | $5$ | $\left\|\begin{array}{\|cc\|} \hline & 0 \\ 0 & 0 \\ -0 & 0 \end{array}\right\|$ |  |  | $0$ |  |  | $\left\|\begin{array}{ll} 0 & 0 \\ 0 & -0 \\ 0 & 0 \end{array}\right\|$ | $\left\|\begin{array}{ll} m & 5 \\ 0 & 0 \end{array}\right\|$ | $\left[\begin{array}{cc} M_{2} & 0 \\ 0 & \vdots \\ 0 & 0 \end{array}\right]$ |
|  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
|  |  |  | $\begin{array}{\|l\|l} 3 \\ 3_{0} \\ 0_{1} \\ 0_{2} \\ \hline \end{array}$ | $\begin{aligned} & -1 \\ & \mathbf{o}^{\prime} \\ & \stackrel{\rightharpoonup}{2} \end{aligned}$ | $\begin{array}{\|l\|l} \hline \frac{\square}{0} \\ \stackrel{\rightharpoonup}{0} \\ \hline \end{array}$ | $\begin{aligned} & 0 \\ & \stackrel{0}{0} \\ & \stackrel{0}{0} \\ & \hline \end{aligned}$ | $\begin{array}{\|l\|l} \frac{0}{0} \\ \stackrel{\rightharpoonup}{0} \\ \hline \end{array}$ |  |  | $\begin{array}{\|l\|} \hline E_{1} \\ E_{0} \\ \hline \\ \hline \end{array}$ |  |  |  |  |  |  |  | $\begin{aligned} & \text { 苟 } \\ & \stackrel{\rightharpoonup}{3} \\ & \hline \end{aligned}$ |  | $\frac{\stackrel{0}{0}}{\underline{O}}$ |

1) Coefficients in bold are significant at the $5 \%$ level in all periods


Appendix 9. Computers: Summary results of 35 pooled OLS-regressions 1)

Appendix 10．Televisions：Summary results of 18 bimonthly WLS－regressions（expenditure weighted）1）

|  |  |  | 发影 | ${ }_{0}^{\sim}$ | \％ | \％ |  |  |  |  |  |  |  | $\underbrace{\circ}_{n}$ |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  | 尔 | $y_{3}^{2}$ | 웅 ${ }^{\circ}$ | \％ive |  |  |  |  |  |  |  |  |  |
|  | $\mathfrak{b l}_{6}^{6}$ |  |  |  | $8$ | $s_{s}^{5}$ | ox | So | Bide io | Bio six | $\therefore \stackrel{y}{\circ}$ | 解右 |  |  |  |  |
|  |  | $0$ |  | $8$ | oo: | $8$ |  | $8_{5}^{0}$ | Bid |  |  | $\theta^{\circ}$ |  |  |  |  |
|  |  |  | 淢 | $\left.\left.\right\|_{0}\right\|_{0} ^{\underline{0}}$ |  |  |  | 淢 |  | 気菏 |  |  |  |  |  |  |
|  |  |  |  |  | 范 |  |  |  |  | 乎 | $\begin{aligned} & \text { 若 } \\ & \stackrel{y}{5} \\ & \hline \end{aligned}$ |  |  |  |  | 菏 |


| \％ |  | 成筦 |  |
| :---: | :---: | :---: | :---: |
|  | $0$ | ${ }^{\circ}$ |  |
|  |  |  |  |
|  |  |  |  |
|  |  | . |  |
| 粊 |  | $\begin{array}{\|l\|l} \frac{0}{0} \\ \frac{0}{3} \\ \hline \end{array}$ |  |





|  |  | gem | stdev | min | max |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Constante | coefficient | 4，557 | 0，166 | 4，343 | 4，824 |
|  | st．error | 0.049 | 0.005 | 0.041 | 0.055 |
| diameter | coefficient | 0，056 | 0，005 | 0，046 | 0，063 |
|  | st．error | 0.002 | 0.000 | 0.001 | 0.002 |
| dolby1 | coeficicent | 0，038 | 0，027 | －0，011 | 0，098 |
|  | st．error | 0.015 | 0.001 | 0.013 | 0.017 |
| dvd | coefficient | 0，624 | 0，078 | 0，526 | 0，765 |
|  | st．error | 0.092 | 0.031 | 0.071 | 0.154 |
| flatsor | coefficient | 0，263 | 0，024 | 0，205 | 0，296 |
|  | st．error | 0.015 | 0.003 | 0.011 | 0.021 |
| freez1 | coefficient | 0，066 | 0，136 | －0，102 | 0，242 |
|  | st．error | 0.022 | 0，008 | 0.014 | 0.041 |
| freq100 | coefficient | 0，334 | 0，041 | 0，259 | 0，395 |
|  | st．error | 0.014 | 0.001 | 0.011 | 0.015 |
| mitpip1 | coefficient | 0,087 | 0，027 | ${ }^{0,029}$ | 0，130 |
|  | st．error | 0.021 | 0.004 | 0.012 | 0.026 |
| mitpip2 | coefficient | －0，007 | 0，044 | －0，072 | 0，071 |
|  | st error | 0.016 | 0.001 | 0.013 | 0.018 |
| model1 | coefficient | ${ }^{-0,020}$ | 0，028 | ${ }^{-0,072}$ | 0，025 |
|  | st．error | 0.016 | 0.004 | 0.011 | 0,021 |
| nicam1 | coefficient | 0，027 | 0，023 | ${ }^{-0,025}$ | 0，057 |
|  | st，error | 0.012 | 0.000 | 0.011 | 0.013 |
| ntuners | coefficient | 0，278 | 0，015 | 0，264 | 0，299 |
|  | st．error | 0.114 | 0.074 | 0，050 | 0,218 |
| pc＿conn | coefficient | 0，017 | 0，082 | －0，166 | 0，181 |
|  | st．error | 0.158 | 0.096 | 0.037 | 0,334 |
| pip1 | coefficient | 0，039 | 0，067 | ${ }^{-0,051}$ | 0，192 |
|  | st．error | 0，050 | 0.023 | 0，028 | 0，129 |
| portab | coefficient | 0，105 | 0，031 | 0，049 | 0，150 |
|  | st，error | 0.026 | 0.003 | 0.021 | 0.029 |
| remote0 | coefficient | 0，051 | 0，153 | －0，080 | 0，374 |
|  | st．error | 1，034 | 1，097 | 0.099 | 3.258 |
| s＿vhs1 | coefficient | 0，002 | 0，032 | －0，048 | 0，086 |
|  | st．error | 0,011 | 0.001 | 0.010 | 0.012 |
| satune1 | coefficient | －0，122 | 0，171 | ${ }^{-0,421}$ | 0，122 |
|  | st．error | 0，361 | 0，404 | 0.049 | 1.216 |
| scratio2 | coefficient | 0，193 | 0，014 | 0，159 | 0，221 |
|  | st．error | 0.012 | 0.001 | 0.010 | 0.014 |
| scrrype2 | coefficient | 0，057 | 0，021 | 0，023 | 0，094 |
|  | st．error | 0.024 | 0.004 | 0.018 | 0.030 |
| soundo | coefficient | －0，154 | 0，028 | －0，197 | －0，095 |
|  | st．error | 0.023 | 0.003 | 0.018 | 0.027 |
| syst2 | coefficient | －0，015 | 0，018 | ${ }^{-0,042}$ | 0，030 |
|  | st．error | 0.015 | 0.001 | 0.013 | 0.017 |
| syst4 | coeficicient | ${ }^{-0,005}$ | ${ }_{0}^{0,052}$ | －0，091 | 0，110 |
| syst6 | coefficient | 0，097 | 0，087 | ${ }^{-0,034}$ | 0，292 |
|  | st．error | 0,133 | 0.112 | 0.027 | 0.364 |
| syst7 | coefficient | －0，036 | 0，058 | －0，144 | 0，114 |
|  | st．error | 0,037 | 0.017 | 0.023 | 0.094 |
| syst8 | coefficient | 0，025 | 0，142 | －0，277 | 0，264 |
|  | st．error | 0,250 | 0，222 | 0,047 | 0.761 |

[^5]Appendix 11. Refrigerators: Summary results of 18 bimonthly WLS-regressions (expenditure weighted) 1)


| $\left(\begin{array}{c\|c} 0 \\ 0 \\ 0 \\ 0 \\ 0 \end{array}\right)$ | No | On | N0 |
| :---: | :---: | :---: | :---: |
|  |  |  | \% |
| Mn | $5$ |  | (1) |
|  | $0$ |  | $0$ |
|  |  |  |  |
| $\begin{array}{\|c} \frac{y}{2} \\ \frac{0}{5} \\ \hline 0 \end{array}$ |  |  | $\begin{array}{\|l} \frac{n}{0} \\ \frac{0}{7} \\ 0 \\ \hline \end{array}$ |


Appendix 12. Washing machines: Summary results of 18 bimonthly WLS-regressions (expenditure weighted) 1)

|  |  | Average | st dev | min | max |
| :---: | :---: | :---: | :---: | :---: | :---: |
| brand23 | coefficient | -0,410 | 0,101 | -0,535 | -0,232 |
|  | st. error | 0,265 | 0,367 | 0,071 | 1,450 |
| brand24 | coefficient | -0,183 | 0,089 | -0,321 | -0,087 |
|  | st. error | 0,398 | 0,274 | 0,181 | 0,865 |
| brand25 | coefficient | -0,373 | 0,059 | -0,435 | -0,209 |
|  | st. error | 0,048 | 0,017 | 0,028 | 0,114 |
| brand26 | coefficient | -0,476 | 0,059 | -0,590 | -0,422 |
|  | st. error | 0,369 | 0,313 | 0,115 | 1,011 |
| brand27 | coefficient | -0,430 | 0,093 | -0,580 | -0,283 |
|  | st. error | 0,180 | 0,091 | 0,092 | 0,472 |
| brand28 | coefficient | -0,501 | 0,034 | -0,565 | -0,437 |
|  | st. error | 0,024 | 0,005 | 0,017 | 0,034 |
| brand29 | coefficient | -0,052 | 0,052 | -0,128 | 0,058 |
|  | st. error | 0,094 | 0,101 | 0,033 | 0,363 |
| brand30 | coefficient | -0,153 | 0,028 | -0,189 | -0,095 |
|  | st. error | 0,032 | 0,005 | 0,023 | 0,045 |
| brand31 | coefficient | 0,277 | 0,023 | 0,237 | 0,309 |
|  | st. error | 0,011 | 0,001 | 0,008 | 0,012 |
| brand32 | coefficient | -0,319 | 0,051 | -0,478 | -0,231 |
|  | st. error | 0,031 | 0,013 | 0,020 | 0,084 |
| brand33 | coefficient | -0,425 | 0,128 | -0,567 | -0,048 |
|  | st. error | 0,330 | 0,114 | 0,122 | 0,473 |
| brand34 | coefficient | -0,686 | 0,086 | -0,810 | -0,558 |
|  | st. error | 0,170 | 0,067 | 0,089 | 0,311 |
| brand35 | coefficient | -0,547 | 0,084 | -0,665 | -0,404 |
|  | st. error | 0,526 | 0,523 | 0,048 | 1,362 |
| brand36 | coefficient | -0,165 | 0,178 | -0,460 | 0,003 |
|  | st. error | 0,137 | 0,022 | 0,102 | 0,163 |
| brand37 | coefficient | -0,321 | 0,055 | -0,429 | -0,230 |
|  | st. error | 0,033 | 0,011 | 0,023 | 0,059 |
| brand38 | coefficient | -0,020 | 0,031 | -0,066 | 0,037 |
|  | st. error | 0,015 | 0,001 | 0,012 | 0,017 |
| brand39 | coefficient | -0,196 | 0,285 | -0,687 | 0,008 |
|  | st. error | 0,170 | 0,167 | 0,064 | 0,458 |
| brand41 | coefficient | -0,152 | 0,125 | -0,314 | 0,206 |
|  | st. error | 0,328 | 0,243 | 0,130 | 1,054 |
| brand42 | coefficient | -0,282 | 0,030 | -0,341 | -0,203 |
|  | st. error | 0,018 | 0,001 | 0,016 | 0,021 |
| brand43 | coefficient | -0,051 | 0,049 | -0,159 | 0,024 |
|  | st. error | 0,029 | 0,004 | 0,023 | 0,039 |
| brand44 | coefficient | -0,263 | 0,023 | -0,308 | -0,222 |
|  | st. error | 0,018 | 0,002 | 0,015 | 0,024 |
| brand45 | coefficient | -0,514 | 0,000 | -0,514 | -0,514 |
|  | st. error | 0,110 | 0,000 | 0,110 | 0,110 |


|  |  |  | $\left\|\begin{array}{cc} \infty \\ \infty & \infty \\ 0 & 0 \\ 0 & 0 \end{array}\right\|$ |  |  | $\dot{\|c c\|}$ |  | $\cdots 5$ | $0$ | $=\left[\begin{array}{ll} 5 & 0 \\ \hline-0 & 0 \end{array}\right.$ | $\begin{array}{\|cc\|} \hline 0 & 8 \\ 0 & -1 \\ -0 & 0 \end{array}$ |  | $0$ |  | $\begin{array}{lll} \hline 0 & 0 \\ 0 \\ 0 & 0 \\ i & 0 \\ \hline \end{array}$ | $\begin{array}{\|l\|l\|} \hline & N_{0} \\ \hline \end{array}$ |  | $0$ |  | $\left\lvert\, \begin{array}{cc} c_{1} & \infty \\ 0 & \infty \\ \hline \end{array}\right.$ |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $\left.\begin{array}{\|cc\|} \hline \begin{array}{c} m \\ \hline \end{array} & 0 \\ 0 & 0 \\ \hline & 0 \\ \hline \end{array} \right\rvert\,$ | $\begin{array}{\|cc\|} \hline \infty & 0 \\ \vdots \\ \hline \end{array}$ | $\begin{array}{\|cc\|} \hline & 0 \\ N_{0} & 0 \\ 0 \end{array}$ |  |  | $0$ | $\begin{array}{ll} \hline 1 & 0 \\ 0 & 0 \\ 0 & 0 \\ 0 \end{array}$ | $\mathfrak{m}$ | $\begin{array}{\|cc\|} \hline 1 & e_{1} \\ 0 & 0 \\ 0 & 0 \end{array}$ | $5$ | $\left.\begin{array}{\|cc\|} \hline-0 & 0 \\ \hdashline-0 \\ \hline-0 & 0 \end{array} \right\rvert\,$ | $5$ | bio |  | $\begin{array}{ll} \infty \\ \hline \end{array}$ |  | $\begin{array}{lll} \infty & 0 \\ \hline \end{array}$ | $\mathfrak{l l}$ |  |  | $1 \begin{array}{ll} 0 \\ 0 & 0 \\ 0 & 0 \\ 0 & 0 \end{array}$ |
| $\stackrel{\rightharpoonup}{2}$ | Non | $\begin{array}{ll} 10 \\ 0 & 0 \\ 0 & 0 \\ 0 & 0 \end{array}$ |  | $\left.\begin{array}{\|cc\|} \hline 20 & 0 \\ 0 & 8 \\ 0 \end{array} \right\rvert\,$ | $\begin{array}{\|c\|c\|} \hline 0 & 0 \\ \hline 0.0 \\ 0-8 \\ 0 \end{array}$ |  |  |  | $\frac{1}{2}$ | for | $\begin{array}{ll} \infty \\ \hline \end{array}$ | $0$ | Si | $\mathfrak{c}$ | $\begin{gathered} m \\ \hline \end{gathered}$ | $0$ |  |  | $\begin{array}{ll} \hline & 0 \\ \hdashline- & 0 \\ 0 & 0 \end{array}$ |  | $\begin{array}{\|cc\|} \hline 0 & \infty \\ 0 & 0 \\ 0 & 0 \\ 0 & 0 \end{array}$ |
| $\stackrel{8}{0}$ |  | $\begin{array}{\|cc\|} \hline y_{0} & 0 \\ \hline \end{array}$ | $$ |  | $\left.\begin{array}{\|cc\|} \hline A_{1} & 0 \\ -i_{0} & 0 \end{array} \right\rvert\,$ |  | $$ | $\begin{array}{\|cc\|} \hline n & 0 \\ \hline & 0 \\ 0 & 0 \\ \hline \end{array}$ | $\begin{array}{\|cc\|} \hline \mathbb{F}_{2} & \text { n } \\ 0 \\ 0 & \infty \\ i & 0 \end{array}$ | $5$ |  | $5$ | $\mathfrak{c c}$ |  | $\begin{array}{ll} \hline \text { On } & 0 \\ \hline \end{array}$ |  |  | \|ros | Oi |  | $\left\lvert\, \begin{array}{ll} \infty & 2 \\ \hline \end{array}\right.$ |
|  |  |  |  |  | $\left\lvert\, \begin{array}{ll} \stackrel{\rightharpoonup}{\mathbf{L}} \\ \hline \end{array}\right.$ |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
|  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |



1) Coefficients in bold are significant at the $5 \%$ level in all periods

| $\stackrel{\times}{\text { ® }}$ | （1） | N | － | $\left\lvert\, \begin{array}{cc} \infty & \underset{n}{c} \\ \\ \hline \end{array}\right.$ | $\left\lvert\, \begin{gathered} 1 \\ \\ \\ \hline \end{gathered}\right.$ | $\left\|\begin{array}{cc} c_{0} \\ \\ 0 & 0 \\ 0 \end{array}\right\|$ | $\left\lvert\, \begin{array}{cc} \substack{0 \\ 0 \\ \vdots \\ 0 \\ 0 \\ 0 \\ \hline} \\ \hline \end{array}\right.$ |  | $\left\|\begin{array}{cc} \hat{N} & \stackrel{\rightharpoonup}{0} \\ & 0 \\ 0 \end{array}\right\|$ | \％ | $\left\|\begin{array}{cc} \hat{y} \\ 0 & y \\ 0 & 0 \\ 0 \end{array}\right\|$ | $\left\|\begin{array}{cc} n & n \\ 0 & N \\ 0 & \underset{y}{n} \end{array}\right\|$ | Nr | No | $\mathfrak{c c}_{\substack{0}}^{0}$ |  |  |  | $\left\|\begin{array}{cc} 0 & 0 \\ \hline & 0 \\ \vdots & 0 \end{array}\right\|$ | $0$ | $\mathfrak{A c}$ | 资 | $\mathfrak{c c}$ | N |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| ¢ | $=\begin{array}{cc} \substack{\circ \\ C} \\ 0 & 0 \\ 0 \end{array}$ | $\left\lvert\, \begin{array}{cc} \tilde{N}_{0} \\ \\ \text { on } \\ \hline \end{array}\right.$ | $\left\lvert\, \begin{array}{cc} \substack{0 \\ 0 \\ 0 \\ 0 \\ \hline \\ \hline \\ \hline} \end{array}\right.$ |  |  |  | $\left\lvert\, \begin{array}{ll} \infty & 0 \\ & 0 \\ & 0 \\ 0 \end{array}\right.$ | $\left\|\begin{array}{cc} \substack { 2 \\ \\ \begin{subarray}{c}{0{ 2 \\ \\ \begin{subarray} { c } { 0 } } \\ {\hline} \end{array}\right\|$ | $\left\|\begin{array}{cc} \infty & \tilde{0} \\ 0 & \tilde{N} \\ 0 & 0 \end{array}\right\|$ | $\left\|\begin{array}{ll} 1 & 0 \\ \text { y } \\ 0 & 0 \\ 0 & 0 \end{array}\right\|$ | $\begin{array}{ll}  & 0 \\ & 0 \\ 0 & 0 \end{array}$ |  | 症 |  | $\begin{array}{ll} \mathrm{N} & \mathrm{~N} \\ \hdashline \mathbf{O} & 0 \\ 0 \end{array}$ | $\left(\begin{array}{cc} 0 \\ \\ \\ 0 & 0 \\ 0 \end{array}\right.$ | $\left\|\begin{array}{cc} \bar{x} & 0 \\ & 0 \\ \\ \hline \end{array}\right\|$ | $\begin{array}{\|cc\|} \infty & \tilde{\sim} \\ \\ 0 \\ 0 \end{array}$ | $\left\|\begin{array}{ll} 10 & 0 \\ \hline 8 & 0 \\ 0 & 0 \end{array}\right\|$ |  | $0$ |  | $\left\lvert\, \begin{array}{cc} ﹎{0}^{2} & 0 \\ 0 & \vdots \\ \hline \end{array}\right.$ | $\left\|\begin{array}{ll} \bar{y} & 0 \\ \hdashline \mathbf{c} \\ \hline 0 & 0 \end{array}\right\|$ |
| $\left\|\begin{array}{c} \frac{3}{0} \\ \stackrel{\rightharpoonup}{i} \end{array}\right\|$ | $0$ | $\left\|\begin{array}{ll} N & 7 \\ \cline { 1 - 3 } & 0 \\ 0 & 0 \end{array}\right\|$ | $\left\|\begin{array}{cc} 0 & 0 \\ \hdashline-0 \\ 0 & 0 \\ 0 \end{array}\right\|$ | $\left\lvert\, \begin{array}{cc} 0 \\ \hline \end{array}\right.$ | $\left\|\begin{array}{ll} 1 & 0 \\ 0 & 0 \\ 0 & 0 \\ 0 \end{array}\right\|$ | $\left\|\begin{array}{cc} x_{0} & 0 \\ 0 & 0 \\ 0 & 0 \end{array}\right\|$ | $\left\|\begin{array}{cc} \infty \\ \frac{1}{4} & \frac{0}{2} \\ 0 \end{array}\right\|$ | $\left\|\begin{array}{cc} N_{2} & 0 \\ 0 & 0 \\ 0 & 0 \end{array}\right\|$ | $\left\|\begin{array}{ll} 0 & 0 \\ 0 & 0 \\ 0 & 0 \\ 0 \end{array}\right\|$ | $\left\lvert\, \begin{array}{cc} \tilde{N}_{2} \\ \text { à } \\ \\ \hline \end{array}\right.$ | $0$ | $\left\lvert\,\right.$ | $\left\|\begin{array}{cc} 0 & \hat{0} \\ 0 & 0 \\ 0 & 0 \\ 0 & 0 \end{array}\right\|$ | $\left[\begin{array}{cc} 0 \\ \hline \end{array}\right.$ | $0$ | $\left\|\right\|$ | $\left\|\begin{array}{cc} 0 \\ \hline & 0 \\ \vdots \\ 0 & 0 \\ \hline \end{array}\right\|$ | $\left\lvert\, \begin{array}{ll\|} \hline 0 & 0 \\ 0 & 0 \\ 0 & 0 \\ \hline \end{array}\right.$ | $\left\|\begin{array}{ll} 9 & 8 \\ 0 & 8 \\ 0 & 0 \end{array}\right\|$ | $0$ | $0$ | $\left\lvert\, \begin{array}{ll} \circ & 0 \\ \hline & 8 \\ 0 & 0 \\ \hline \end{array}\right.$ | $0$ | $\left\|\begin{array}{ll} \hat{A} & 0 \\ 0 & j \\ 0 & 0 \end{array}\right\|$ |
|  |  | $\left\|\begin{array}{cc} \tilde{N}_{1} & \tilde{y} \\ 0 \\ 0 \end{array}\right\|$ | $\left\|\begin{array}{cc} \substack{m \\ 0 \\ 0 \\ 0 \\ 0} & 0 \\ 0 \end{array}\right\|$ |  | $\left\lvert\, \begin{array}{cc} \infty & \approx \\ 8 \\ \hline & 0 \\ 0 \end{array}\right.$ | $\left\|\begin{array}{ll} 10 & 0 \\ \hdashline & 0 \\ 0 & 0 \end{array}\right\|$ | $\left\|\begin{array}{ll} 0 & 2 \\ 0 & 2 \\ & \stackrel{0}{0} \end{array}\right\|$ | $\left\lvert\, \begin{array}{\|cc\|} \hline 0 & 0 \\ \\ \\ \hline \end{array}\right.$ | $\left\|\begin{array}{cc} n & 0 \\ 0 & 0 \\ 0 & 0 \\ 0 & 0 \end{array}\right\|$ | $\left\|\begin{array}{cc} 0 & 0 \\ & 0 \\ \hdashline-0 & 0 \\ 0 & 0 \end{array}\right\|$ |  | $\begin{array}{ll} 1 & 0 \\ \hline & 0 \\ 0 \\ 0 & 0 \end{array}$ | Ur | $\mathfrak{c c}$ | $0$ | $\left\lvert\, \begin{array}{cc} 10 & 0 \\ 0 & 0 \\ 0 & 0 \end{array}\right.$ | $\left\lvert\, \begin{array}{ll} \infty \\ \hline \end{array}\right.$ | $\left\lvert\, \begin{array}{cc} \infty \\ \\ \hline \end{array}\right.$ | $\left\|\begin{array}{cc} \begin{array}{c} 0 \\ 0 \\ 0 \end{array} & 0 \\ 0 & 0 \end{array}\right\|$ | $\left[\begin{array}{ll} \infty & \infty \\ 0 & \vdots \\ 0 & 0 \\ 0 \end{array}\right.$ | $0$ | $\left\|\begin{array}{cc} 0_{0}^{0} \\ \text { on } \\ 0 & 0 \end{array}\right\|$ | $\left\|\begin{array}{cc} \overline{8} & \tilde{0} \\ -0 & 0 \end{array}\right\|$ |  |
|  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
|  | $\begin{aligned} & \stackrel{N}{\stackrel{N}{c}} \\ & \stackrel{y}{0} \end{aligned}$ |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  | $\begin{aligned} & \text { 号 } \\ & \text { 合 } \\ & \hline \end{aligned}$ |  |  | $\stackrel{\stackrel{\rightharpoonup}{0}}{\underline{c}}$ | 흘 | \％ |

Appendix 13．Computers：Summary results of 35 pooled WLS－regressions（Diewert－variant）1）

| ¢ $\times$ ¢ | 욱웅 | － | －iocro | － | － |  |  |  | $\left\lvert\, \begin{array}{ll} 8 & 20 \\ 0 \\ 0 & 4 \\ \hline \end{array}\right.$ | $\left\|\begin{array}{ll} n & 0 \\ \\ 0 & 0 \\ 0 & 0 \end{array}\right\|$ | $\underbrace{\infty}_{0}$ | $\mathfrak{c c}$ | $\begin{array}{ll} \infty \\ \infty \\ \infty \\ 0 \\ \hline \end{array}$ | $\left\|\begin{array}{cc} Z_{2} & 0 \\ 0 & 0 \\ 0 & 0 \end{array}\right\|$ | $\mathfrak{c c}$ | $\left\|\begin{array}{ll} \infty \\ 0 & 0 \\ 0 \\ 0 & 0 \\ i \end{array}\right\|$ | $\left\lvert\, \begin{array}{cc} 0 & 0 \\ & 8 \\ 0 & 0 \\ 0 \end{array}\right.$ | $0$ |  | $\left\|\begin{array}{cc} 1 & \ddots \\ 0 & 8 \\ 0 & 0 \\ 0 \end{array}\right\|$ | $\left\|\begin{array}{cc} N_{0}^{\circ} & 0 \\ 0-0 \\ 0 & 0 \end{array}\right\|$ | $\left\|\begin{array}{cc} \infty \\ 0 & 0 \\ 0 \\ 0 & 0 \\ 0 & 0 \end{array}\right\|$ |  | $0$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $\left(\left.\begin{array}{cc} \bar{c} & \begin{array}{c} \infty \\ \infty \\ 0 \\ 0 \end{array} \\ \vdots & 0 \end{array} \right\rvert\,\right.$ | $\begin{cases}0 & 0 \\ \\ 0 \\ 0 \\ 0\end{cases}$ |  | $\left\|\begin{array}{cc} N_{0} & 1 \\ 0 & 0 \\ 0 & 0 \end{array}\right\|$ | $\left\|\begin{array}{cc} \infty \\ 0 & \underset{e}{0} \\ 0 & 0 \\ \hline \end{array}\right\|$ |  |  | $\left\|\begin{array}{ccc} n_{2} & \hat{0} \\ 0 & 0 \\ 0 & 0 \end{array}\right\|$ | $\left\|\begin{array}{cc} \hat{N} & \mathbf{0} \\ 0 \\ 0 & 0 \\ 0 \end{array}\right\|$ | $\left\|\right\|$ |  |  | $\mathfrak{c}$ | $\begin{array}{ll} \infty \\ 0 \\ 0 \\ 0 \\ \hline \end{array}$ | $\left\{\begin{array}{cc} w_{0} & 0 \\ 0 & \stackrel{y}{0} \\ \hline \end{array}\right.$ |  |  | $\begin{cases}\infty & \multicolumn{1}{c}{0} \\ 0 & 0 \\ 0 & 0 \\ 0\end{cases}$ | $\left\|\begin{array}{cc} O_{0} & 0 \\ 0 \\ 0 & 0 \\ & 0 \end{array}\right\|$ |  | $\left\lvert\, \begin{array}{cc} \substack{N \\ \\ \vdots \\ 0} \\ \hline \end{array}\right.$ | $\left\|\begin{array}{cc} n_{n}^{N} \\ \\ & \hat{b} \\ \hline \end{array}\right\|$ |  | $\left\lvert\, \begin{array}{cc} \tilde{N}_{2} & 0 \\ 0 & 0 \\ \hline-0 & 0 \end{array}\right.$ | On |
|  | $0$ | $\left\|\begin{array}{cc} \infty & 0 \\ \hline & \tilde{y} \\ 0 & 0 \\ 0 \end{array}\right\|$ | $\left\lvert\, \begin{array}{cc} \infty & 8 \\ 0 \\ \vdots & 8 \\ 0 & 0 \\ 0 \end{array}\right.$ | $\left\|\begin{array}{ll} 0 & \tilde{0} \\ \hline & \tilde{z} \\ 0 & 0 \end{array}\right\|$ | $\begin{array}{ll} 0 & 8 \\ & 0 \\ -1 & 0 \end{array}$ | $\begin{array}{\|cc\|} \substack{2 \\ \hline \\ 0 \\ 0} & 0 \\ 0 \end{array}$ |  | $\left\|\begin{array}{cc} \bar{y} & \tilde{y} \\ 0 & 0 \\ 0 \end{array}\right\|$ | $\left\|\begin{array}{cc} 8 & 0 \\ \\ 0 & 0 \\ 0 \end{array}\right\|$ | $\begin{array}{ll} \mathrm{A}_{1} & \infty \\ \mathrm{~N}_{2} & \vdots \\ 0 & 0 \\ \hline \end{array}$ | Nr |  | $\left\lvert\, \begin{array}{cc} 8 & 8 \\ 0 & 0 \\ 0 \end{array}\right.$ |  | $:$ | $\left\|\begin{array}{cc} O_{0}^{\circ} & 0 \\ 0 & 0 \\ 0 \end{array}\right\|$ | $\left\|\begin{array}{cc} 0 \\ 0 & \tilde{0} \\ 0 & 0 \\ 0 \end{array}\right\|$ | $0$ |  | $\left\lvert\, \begin{array}{cc} 8 & 0 \\ \hline-2 & 0 \\ -0 \end{array}\right.$ | $\left\|\begin{array}{ll} 0 & 0 \\ 0 & 0 \\ 0 & 0 \\ 0 & 0 \end{array}\right\|$ | $\left\|\begin{array}{ll} \infty & 0 \\ 0 & 0 \\ \\ 0 \end{array}\right\|$ | $\begin{array}{ll} 0 & 0 \\ 0 & \vdots \\ 0 & 0 \\ 0 \end{array}$ | $\begin{array}{ll} 0 & 0 \\ 0 & 0 \\ 0 \\ 0 & 0 \end{array}$ |
|  | $0 \begin{gathered} 9 \\ 0 \\ 0 \\ 0 \\ 0 \end{gathered}$ |  | $\left\lvert\, \begin{array}{lll} \infty & 0 \\ y & 0 \\ 0 & 8 \\ 0 \end{array}\right.$ | $\left\|\right\|$ | $\left\|\begin{array}{cc} \circ & 0 \\ & 0 \\ -0 \\ -0 & 0 \end{array}\right\|$ | $\left\lvert\, \begin{array}{cc\|} \hline-0 & 0 \\ 0 & 0 \\ 0 & 0 \end{array}\right.$ | $\left\|\begin{array}{cc} 1 & 2 \\ 0 & \underset{3}{0} \\ 0 & 0 \end{array}\right\|$ | $\left\|\begin{array}{ll} \infty & 0 \\ ⿵ 人 一 ⿰ ⺝ 刂 \\ \hline \end{array}\right\|$ | $\left\|\begin{array}{ll} 4 & 0 \\ 0 & 0 \\ 0 & 0 \\ 0 & 0 \end{array}\right\|$ |  | $\begin{array}{cc} 1 & 0 \\ -0 & 8 \\ 0 & 0 \end{array}$ | $3 \begin{array}{ll} N & 0 \\ & 2 \\ 0 & 0 \\ 0 \end{array}$ | $\left\lvert\, \begin{array}{cc} \infty & 0 \\ 0 & 8 \\ -1 \\ \end{array}\right.$ | $0$ |  |  | $\left\|\begin{array}{cc} 1 \\ \hline 0 \\ 0 \\ 0 & 0 \\ 0 & 0 \end{array}\right\|$ | $\left\|\begin{array}{cc} 10 \\ 0 \\ 0 \\ 0 & z \\ i \end{array}\right\|$ | No |  |  | $\left\|\begin{array}{cc} 0 & \hat{0} \\ 0 & \hat{0} \\ 0 & 0 \end{array}\right\|$ | $\left\|\begin{array}{cc} y_{0} & 0 \\ 0 & 0 \\ 0 & 0 \end{array}\right\|$ |  |
|  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
|  | $\begin{array}{\|l\|l\|} \hline \stackrel{\text { O}}{2} \\ \text { O} \end{array}$ | $\begin{array}{\|l} \overline{\mathrm{O}} \\ \text { oin } \end{array}$ | $\begin{array}{\|l\|} \substack{\mathrm{N} \\ \mathrm{o} \\ \hline} \end{array}$ |  | $$ | \| |  | $\begin{array}{\|c} \hat{\mathrm{O}} \\ \text { ob } \end{array}$ |  |  |  | $\begin{aligned} & \text { 䓂 } \\ & \frac{0}{5} \\ & \frac{0}{0} \end{aligned}$ | $\left\lvert\, \begin{aligned} & \stackrel{0}{0} \\ & \stackrel{0}{C} \\ & \stackrel{i}{0} \\ & \hline \end{aligned}\right.$ |  |  |  |  | $\begin{array}{\|l} \frac{0}{0} \\ \frac{0}{0} \\ \text { Wig } \\ \hline \end{array}$ |  | $\begin{array}{\|l\|l\|} \hline \stackrel{y}{\mathrm{~N}} \\ \stackrel{\rightharpoonup}{\mathrm{y}} \\ \hline \end{array}$ |  |  |  |  |


|  | － | $\left\|\begin{array}{ll} 0 & 0 \\ 0 & 0 \\ 0 & 0 \end{array}\right\|$ | $\left\lvert\, \begin{array}{ll} 10 & 8 \\ 0 & 8 \\ 0 & 0 \\ 0 \end{array}\right.$ | $\begin{array}{ll} 0 & 8 \\ 0 & 8 \\ 0 & 0 \\ 0 \end{array}$ | $\left\|\begin{array}{ll} n & \underset{N}{n} \\ 0 & \lambda \\ 0 & 0 \\ 0 & 0 \end{array}\right\|$ | $\left\|\begin{array}{cc} \substack{n \\ N} \\ 0 & N \\ 0 & -1 \end{array}\right\|$ | $\begin{array}{ll} 1 & 0 \\ 0 & 0 \\ 0 & 0 \\ 0 & -1 \end{array}$ |  |  | $\left\|\begin{array}{ll} 0 & 0 \\ & 0 \\ 0 & 0 \\ 0 \end{array}\right\|$ | $\mathfrak{c}$ | $\left\lvert\, \begin{array}{ll} N_{2} & \mathbb{N} \\ 0 & 0 \\ 0 & 0 \end{array}\right.$ | $\begin{array}{ll} m & \circ \\ 50 & 9 \\ 0 & 0 \\ 0 \end{array}$ | $\left\|\begin{array}{cc} 0 & \infty \\ 0 & \infty \\ 0 & 0 \\ 1 & 0 \\ 0 \end{array}\right\|$ | $\left\lvert\, \begin{array}{cc} x_{0}^{\infty} & 0 \\ 0 & 0 \\ 0 & 0 \end{array}\right.$ | $\left\|\begin{array}{cc} \bar{N} & 0 \\ 0 & \vdots \\ 0 & 0 \end{array}\right\|$ | $\left\lvert\,\right.$ | $\left\|\begin{array}{cc} 0 & 0 \\ 0 & \infty \\ 0 & \infty \\ 0 \end{array}\right\|$ | $\left\|\begin{array}{ll} \infty & 0 \\ \infty & \infty \\ \\ 0 & - \\ 0 \end{array}\right\|$ |  |  |  | $\left\lvert\, \begin{array}{ll} \infty & \mathbb{N} \\ & \underset{\sim}{0} \\ \hline \end{array}\right.$ | $\left\|\begin{array}{cc} \underset{\sim}{\infty} & \infty \\ \underset{\sim}{c} & \infty \\ - \end{array}\right\|$ | $\begin{array}{ll} \hat{N} & \hat{o} \\ \frac{1}{2} & \frac{0}{0} \end{array}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| － | $\left\|\right\|$ | $\left\|\begin{array}{ll} 0 & 8 \\ 0 & 8 \\ 0 & 0 \end{array}\right\|$ | $1 \begin{array}{ll} 8 & 8 \\ 0 & 8 \\ 0 & 0 \end{array}$ | $\left\|\begin{array}{ll} O & 8 \\ \hline & 8 \\ 0 & 0 \end{array}\right\|$ | $\left\|\begin{array}{ll} 0 & \infty \\ 0 & 5 \\ 0 & 0 \\ 0 & 0 \end{array}\right\|$ |  | $\left\|\begin{array}{cc} 0 & 0 \\ 0 & 0 \\ & 0 \\ -1 \end{array}\right\|$ | $\left\lvert\, \begin{array}{ll} N & 8 \\ & 8 \\ - & -1 \end{array}\right.$ | $\left\lvert\, \begin{array}{ll} n & d \\ 0 & 0 \\ 0 & 0 \\ 0 & 0 \end{array}\right.$ | $\left\|\begin{array}{cc} H \\ m & 0 \\ \hdashline- & 0 \\ i & 0 \end{array}\right\|$ | $\left\|\begin{array}{ll} 0 & 0 \\ 0 & 0 \\ & \vdots \\ 1 & 0 \end{array}\right\|$ | $\left\lvert\, \begin{array}{cc} \tilde{N}_{2} & \bar{y} \\ & 0 \\ 1 \end{array}\right.$ | $\left\lvert\, \begin{array}{ll} \infty & 0 \\ 0 & \vdots \\ 0 & 0 \\ 1 & 0 \end{array}\right.$ |  | $\begin{array}{cc} m & 0 \\ \text { N} & 0 \\ 0 & 0 \end{array}$ | $\left\|\begin{array}{ll} n \\ \infty & \hat{0} \\ 0 & 0 \\ 0 & 0 \end{array}\right\|$ |  | $\left\|\begin{array}{cc} 1 & \infty \\ 0 & \infty \\ 0 & \infty \\ 0 \\ 0 & 0 \end{array}\right\|$ | $\left\|\begin{array}{ll} \bar{N} & 0 \\ \mathbf{N}_{2} & \vdots \\ i & 0 \end{array}\right\|$ | $\left\lvert\, \begin{array}{ll} 4 & 0 \\ \infty & 0 \\ \hdashline i & 0 \\ 1 \end{array}\right.$ | $\left(\begin{array}{ll} \infty & \underset{y y}{*} \\ 0 & 0 \\ 0 & 0 \\ \hline \end{array}\right.$ | 会产 | $\left\lvert\, \begin{array}{ll} 0 & 1 \\ \hat{n} & 0 \\ 0 & 0 \\ 1 & 0 \end{array}\right.$ |  | M |
| ¢ | $\left\|\begin{array}{cc} 4 \\ \infty \\ \vdots & 0 \\ 0 & 0 \\ 0 & 0 \end{array}\right\|$ | $\left\|\begin{array}{ll} \overline{5} & 8 \\ 0 & 0 \\ 0 & 0 \end{array}\right\|$ | $\left\|\begin{array}{cc} 8 & 8 \\ 0 & 8 \\ 0 & 0 \end{array}\right\|$ | $\left\|\begin{array}{ll} 5 & 8 \\ 0 & 0 \\ 0 & 0 \end{array}\right\|$ | $\left\|\begin{array}{ll} 2 & 0 \\ 0 & 0 \\ 0 & 0 \end{array}\right\|$ | $\left\lvert\, \begin{array}{ll} 1 & 0 \\ 0 & 0 \\ -0 & 0 \end{array}\right.$ | $\left\|\begin{array}{cc} n & \hat{n} \\ & \underset{y}{c} \\ 0 & 0 \end{array}\right\|$ | $\left\|\begin{array}{ll} \frac{1}{2} & \overline{8} \\ 0 & 0 \end{array}\right\|$ | $\left(\left.\begin{array}{cc} 1 & 0 \\ & 0 \\ 0 & 0 \\ 0 \end{array} \right\rvert\,\right.$ | $\left\|\begin{array}{ll} 0 & N \\ 0 & \tilde{0} \\ 0 & 0 \end{array}\right\|$ | $\begin{array}{ll} 0 & 0 \\ 0 & 8 \\ 0 & 0 \end{array}$ | $\left\lvert\, \begin{array}{ll} \infty & \infty \\ 0 & \vdots \\ 0 & \vdots \\ 0 & 0 \end{array}\right.$ | $\left[\begin{array}{ll} 0 & \tilde{0} \\ & 0 \\ 0 & 0 \end{array}\right.$ | $\hat{\mid c c}$ | $\left\|\begin{array}{ll} \infty & 0 \\ 0 & N \\ 0 & 0 \end{array}\right\|$ |  | $\left\lvert\, \begin{array}{cc} \bar{j} & \underset{y}{2} \\ \underset{\sim}{\circ} & 0 \end{array}\right.$ | $\left\|\begin{array}{cc} 0 & \infty \\ 0 & 0 \\ 0 & 0 \\ 0 & 0 \end{array}\right\|$ | $\left\|\right\|$ | $\left\lvert\, \begin{array}{ll} \bar{o} & \hat{o} \\ \hdashline-1 & 0 \end{array}\right.$ | $\left\|\begin{array}{ll} \infty & 6 \\ \hdashline & 0 \\ 0 & 0 \\ 0 \end{array}\right\|$ | $\left\lvert\, \begin{array}{cc} \infty & 0 \\ 0 & \infty \\ 0 & \infty \\ 0 & 0 \\ \hline \end{array}\right.$ | $\left\|\begin{array}{ll} \infty & N \\ & \underset{\sim}{0} \\ \hline \end{array}\right\|$ | $\left\|\begin{array}{\|cc\|} \hline 0 & 7 \\ 0 & 0 \\ 0 & 0 \\ 0 & 0 \end{array}\right\|$ | Nr |
| $\left\|\begin{array}{c} \stackrel{\otimes}{0} \\ \stackrel{\rightharpoonup}{0} \\ \stackrel{\rightharpoonup}{\sigma} \end{array}\right\|$ | $\left.\begin{array}{ll} 0 & - \\ 0 & \underset{0}{0} \\ 0 & 0 \end{array} \right\rvert\,$ | $\left\|\begin{array}{cc} \substack{0 \\ \underset{\sim}{w}} & 0 \\ \underset{\sim}{\infty} & 0 \\ \sim \end{array}\right\|$ |  |  | $\left\|\begin{array}{cc} N_{0} & - \\ \mathcal{M} & \mathbf{O} \end{array}\right\|$ | $\left\lvert\, \begin{array}{cc} \mathfrak{y} & 0 \\ y_{0} & 0 \\ 0 \end{array}\right.$ | $\begin{cases}N & 0 \\ & \underset{y}{c} \\ \hline\end{cases}$ | － | $\left(\begin{array}{ll} \mathbf{c}_{1} & 0 \\ \infty & 0 \\ 0 & 0 \end{array}\right.$ | $\left\|\begin{array}{ll} m & 2 \\ \vdots & \tilde{0} \\ & 0 \end{array}\right\|$ | $\mathfrak{c}$ | $\left\lvert\, \begin{array}{cc} n \\ \hdashline \sigma_{i} & \hat{Z} \end{array}\right.$ | $\left[\begin{array}{ll} 0 & \hat{0} \\ 0 & 0 \\ 0 & 0 \end{array}\right.$ | $\mathfrak{l}$ | $\left\lvert\, \begin{array}{ll} \bar{F} & \bar{c} \\ \hdashline i & 0 \end{array}\right.$ | $\left\|\begin{array}{cc} \bar{N} & \hat{N} \\ \mathbf{N}_{2} & \mathbf{N} \end{array}\right\|$ | $\left\lvert\, \begin{array}{ll} N & \hat{N} \\ & \hat{0} \\ i & 0 \end{array}\right.$ | $\left\|\begin{array}{cc} N_{N} & \overline{0} \\ & 0 \\ i & 0 \end{array}\right\|$ | $\left\|\begin{array}{ll} \hat{y} & \hat{y} \\ \\ i & 0 \end{array}\right\|$ | $\left\|\begin{array}{ll} 0 & \hat{0} \\ 0 & 0 \\ 0 & 0 \end{array}\right\|$ | $\left\|\begin{array}{cc} 0 & \infty \\ 0 & 0 \\ 0 & \tilde{y} \\ \hline & 0 \end{array}\right\|$ | $\left\lvert\, \begin{array}{ll} 0 & \\ 0 & 0 \\ 0 & 0 \\ 0 & 0 \\ 1 & 0 \end{array}\right.$ | $\left\|\begin{array}{cc} \bar{\infty} & 0 \\ & 0 \\ \hline-1 & 0 \end{array}\right\|$ | $\left\|\right\|$ | $\left\lvert\, \begin{array}{cc} 0_{0} & 0 \\ 0 \\ 0 & 0 \\ 0 \end{array}\right.$ |
|  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
|  |  |  |  |  |  |  | － | $\begin{array}{\|l} \stackrel{\rightharpoonup}{0} \\ 0 \\ 0 \\ \hline \end{array}$ | $\begin{aligned} & \text { No } \\ & \text { O} \\ & \hline 0 \end{aligned}$ | $\begin{aligned} & \text { O} \\ & \text { O} \\ & \text { O} \end{aligned}$ | $\begin{aligned} & \text { T } \\ & 0 \\ & 0 \\ & 0 \end{aligned}$ | $\begin{aligned} & \text { OO } \\ & 0 \\ & 0 \\ & 0 \\ & \hline 0 \end{aligned}$ | $\begin{array}{\|l\|l\|} \hline 0 \\ \hline 0 \\ \hline 0 \\ \hline 0.0 \end{array}$ | $\begin{aligned} & \hat{0} \\ & 0 \\ & 0.0 \\ & \hline 0 \end{aligned}$ | $\begin{aligned} & 00 \\ & 00 \\ & 0.0 \\ & \text { on } \end{aligned}$ | $\begin{aligned} & \text { O} \\ & \text { O} \\ & \text { O} \\ & \hline 0 \end{aligned}$ | $$ | $\begin{aligned} & \overline{\mathrm{F}} \\ & \mathrm{O} \\ & \hline \end{aligned}$ | $\begin{aligned} & \text { N } \\ & \stackrel{y}{0} \\ & \hline \end{aligned}$ | $\begin{aligned} & \stackrel{M}{⿺} \\ & \hline \mathrm{O} \\ & \hline \end{aligned}$ | $\begin{aligned} & \stackrel{\rightharpoonup}{\%} \\ & \frac{0}{2} \end{aligned}$ | $$ | $\begin{aligned} & \circ \\ & \stackrel{\circ}{0} \\ & \hline \frac{0}{2} \end{aligned}$ | $\begin{aligned} & \stackrel{\rightharpoonup}{0} \\ & \stackrel{0}{2} \end{aligned}$ | $\begin{aligned} & \frac{\infty}{0} \\ & \hline \frac{0}{2} \\ & \hline \end{aligned}$ |

## Appendix 14.

Price index numbers for televisions (199902=100)

|  | $P_{G T}$ | $P_{G T(i)}$ | $P_{M T}$ | WLS time <br> dummy (exp. <br> shares) | WLS time <br> dummy <br> (quantities) | OLS time <br> dummy | CPI $^{1)}$ |
| :--- | :--- | :--- | :--- | :---: | :---: | :---: | :---: |
| 199902 | 100.00 | 100.00 | 100.00 | 100.00 | 100.00 | 100.00 | 100.0 |
| 199904 | 97.26 | 97.01 | 96.95 | 97.42 | 97.93 | 97.57 | 97.50 |
| 199906 | 94.97 | 94.73 | 94.87 | 95.42 | 96.22 | 95.78 | 96.17 |
| 199908 | 93.24 | 92.99 | 93.25 | 93.84 | 94.32 | 94.42 | 94.54 |
| 199910 | 90.86 | 90.75 | 91.10 | 91.82 | 92.74 | 92.70 | 93.06 |
| 199912 | 89.62 | 89.44 | 89.79 | 91.02 | 91.77 | 90.87 | 91.73 |
| 200002 | 88.43 | 88.24 | 88.72 | 89.90 | 90.61 | 89.94 | 90.92 |
| 200004 | 87.81 | 87.61 | 87.63 | 89.53 | 89.84 | 88.91 | 90.84 |
| 200006 | 87.16 | 87.07 | 87.31 | 88.30 | 88.44 | 97.89 | 90.79 |
| 200008 | 86.44 | 86.34 | 86.63 | 87.82 | 87.96 | 86.90 | 90.04 |
| 200010 | 85.31 | 85.27 | 85.77 | 87.18 | 87.53 | 85.45 | 89.56 |
| 200012 | 85.17 | 85.01 | 85.42 | 87.51 | 87.64 | 85.64 | 88.97 |
| 200102 | 85.55 | 85.34 | 85.71 | 88.13 | 88.25 | 85.22 | 89.27 |
| 200104 | 85.72 | 85.48 | 85.70 | 87.99 | 88.60 | 85.87 | 89.52 |
| 200106 | 85.75 | 85.52 | 85.76 | 88.48 | 89.07 | 85.95 | 89.91 |
| 200108 | 84.90 | 84.56 | 84.77 | 87.75 | 88.31 | 85.57 | 88.28 |
| 200110 | 83.99 | 83.68 | 84.10 | 86.98 | 87.92 | 85.21 | 86.85 |
| 200112 | 83.25 | 82.97 | 83.44 | 86.45 | 87.61 | 85.45 | 86.57 |

${ }^{15}$ Based on official CPI-data.

## Appendix 15.

Price index numbers for refrigerators (199902=100)

|  | $P_{G T}$ | $P_{G T(i)}$ | $P_{M T}$ | WLS time <br> dummy (exp. <br> shares) | WLS time <br> dummy <br> (quantities) | OLS time <br> dummy | CPI ${ }^{1)}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 199902 | 100.00 | 100.00 | 100.00 | 100.00 | 100.00 | 100.00 | 100.00 |
| 199904 | 99.01 | 99.02 | 99.09 | 98.89 | 98.88 | 99.14 | 98.11 |
| 199906 | 97.71 | 97.81 | 98.06 | 97.71 | 97.40 | 97.81 | 98.26 |
| 199908 | 97.37 | 97.51 | 97.68 | 97.76 | 97.29 | 97.98 | 97.84 |
| 199910 | 97.41 | 97.49 | 97.26 | 98.27 | 97.56 | 97.11 | 97.21 |
| 199912 | 96.75 | 96.86 | 96.79 | 98.05 | 97.41 | 97.09 | 97.88 |
| 200002 | 96.12 | 96.23 | 96.26 | 98.06 | 97.06 | 96.04 | 97.86 |
| 200004 | 96.02 | 96.02 | 95.61 | 97.93 | 96.58 | 96.28 | 97.01 |
| 200006 | 95.21 | 95.15 | 94.64 | 97.38 | 95.97 | 95.35 | 98.16 |
| 200008 | 94.29 | 94.24 | 93.77 | 96.28 | 94.92 | 93.58 | 97.55 |
| 200010 | 94.04 | 94.02 | 93.72 | 96.73 | 95.63 | 93.83 | 97.92 |
| 200012 | 94.30 | 94.22 | 93.61 | 97.25 | 95.44 | 94.55 | 98.68 |
| 200102 | 95.16 | 95.09 | 94.27 | 98.45 | 96.10 | 95.93 | 100.49 |
| 200104 | 95.76 | 95.76 | 95.09 | 99.05 | 96.81 | 96.91 | 100.03 |
| 200106 | 96.00 | 95.97 | 94.46 | 99.84 | 96.94 | 96.68 | 101.11 |
| 200108 | 94.72 | 94.8 | 93.57 | 98.28 | 95.45 | 95.99 | 100.95 |
| 200110 | 94.58 | 94.07 | 93.41 | 98.87 | 95.83 | 95.94 | 99.91 |
| 200112 | 94.57 | 94.85 | 93.20 | 100.02 | 96.23 | 95.80 | 99.59 |

${ }^{1)}$ Based on official CPI-data.

## Appendix 16.

Price index numbers for washing machines (199902=100)

|  | $P_{G T}$ | $P_{G T(i)}$ | $P_{M T}$ | WLS time <br> dummy (exp. <br> shares) | WLS time <br> dummy <br> (quantities) | OLS time <br> dummy | CPI ${ }^{1)}$ |
| :--- | :--- | :--- | :--- | :---: | :---: | :---: | :---: |
| 199902 | 100.00 | 100.00 | 100.00 | 100.00 | 100.00 | 100.00 | 100.00 |
| 199904 | 98.03 | 98.01 | 97.96 | 98.39 | 98.06 | 99.16 | 99.27 |
| 199906 | 98.03 | 97.87 | 97.56 | 98.61 | 98.05 | 98.14 | 98.94 |
| 199908 | 98.20 | 98.18 | 97.93 | 99.16 | 98.61 | 97.81 | 99.02 |
| 199910 | 97.76 | 97.73 | 97.54 | 99.07 | 98.33 | 97.76 | 98.04 |
| 199912 | 97.09 | 97.15 | 96.94 | 98.78 | 98.25 | 96.67 | 98.95 |
| 200002 | 96.55 | 96.50 | 96.19 | 97.77 | 96.99 | 96.63 | 97.94 |
| 200004 | 96.07 | 96.00 | 95.66 | 97.75 | 96.99 | 96.32 | 96.83 |
| 200006 | 95.24 | 95.18 | 94.79 | 96.74 | 96.03 | 95.56 | 96.32 |
| 200008 | 93.82 | 93.86 | 93.64 | 95.64 | 94.78 | 94.68 | 95.63 |
| 200010 | 9308 | 93.15 | 93.00 | 94.68 | 93.80 | 93.31 | 97.01 |
| 200012 | 92.66 | 92.69 | 92.47 | 94.49 | 93.57 | 93.43 | 96.91 |
| 200102 | 92.91 | 92.93 | 92.62 | 94.83 | 93.85 | 93.36 | 97.09 |
| 200104 | 93.68 | 93.64 | 93.44 | 96.41 | 95.34 | 94.01 | 98.55 |
| 200106 | 93.59 | 93.46 | 93.03 | 96.26 | 95.00 | 94.33 | 97.76 |
| 200108 | 92.82 | 92.73 | 92.57 | 95.68 | 94.12 | 93.59 | 97.48 |
| 200110 | 92.32 | 92.16 | 92.06 | 95.03 | 93.38 | 92.71 | 96.99 |
| 200112 | 91.58 | 91.39 | 91.21 | 94.75 | 92.89 | 92.25 | 96.70 |

${ }^{1)}$ Based on official CPI-data.

## Appendix 17.

Price index numbers for computers (199901=100)

|  | $P_{G T}$ | $P_{G T(i)}$ | $P_{M T}$ | WLS time dummy (exp. shares) | WLS time dummy (quantities) | OLS time dummy | CPI ${ }^{1)}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 199901 | 100.00 | 100.00 | 100.00 | 100.00 | 100.00 | 100.00 | 100.00 |
| 199902 | 96.28 | 96.02 | 94.77 | 94.87 | 94.50 | 96.55 | 94.42 |
| 199903 | 92.71 | 92.44 | 91.73 | 91.55 | 91.19 | 90.20 | 93.95 |
| 199904 | 88.92 | 88.66 | 89.17 | 87.35 | 87.16 | 88.76 | 90.21 |
| 199905 | 85.68 | 85.70 | 87.05 | 84.02 | 82.91 | 83.59 | 86.73 |
| 199906 | 82.41 | 82.42 | 84.05 | 81.01 | 80.34 | 79.53 | 82.25 |
| 199907 | 77.26 | 77.13 | 78.20 | 77.49 | 76.53 | 77.04 | 75.59 |
| 199908 | 70.80 | 70.68 | 72.51 | 70.56 | 70.42 | 73.66 | 73.98 |
| 199909 | 67.47 | 67.32 | 69.30 | 67.07 | 66.72 | 69.73 | 69.56 |
| 199910 | 67.87 | 67.64 | 69.69 | 67.46 | 67.16 | 70.08 | 68.07 |
| 199911 | 66.31 | 66.29 | 68.27 | 65.88 | 65.78 | 68.46 | 63.66 |
| 199912 | 64.88 | 64.95 | 67.61 | 64.90 | 65.30 | 66.80 | 61.93 |
| 200001 | 64.30 | 64.35 | 66.63 | 63.85 | 63.90 | 65.49 | 60.78 |
| 200002 | 62.88 | 63.24 | 66.37 | 62.02 | 62.33 | 65.98 | 60.17 |
| 200003 | 62.59 | 63.11 | 66.03 | 61.43 | 61.70 | 64.16 | 56.25 |
| 200004 | 61.67 | 62.18 | 65.33 | 60.38 | 60.83 | 63.05 | 56.78 |
| 200005 | 60.31 | 60.87 | 63.92 | 59.09 | 59.29 | 62.82 | 53.01 |
| 200006 | 56.56 | 57.21 | 60.55 | 55.53 | 55.99 | 58.69 | 51.67 |
| 200007 | 55.72 | 56.27 | 59.50 | 55.21 | 55.64 | 56.17 | 48.91 |
| 200008 | 54.37 | 55.09 | 58.59 | 54.01 | 54.57 | 56.84 | 48.05 |
| 200009 | 53.16 | 54.14 | 57.92 | 52.89 | 53.49 | 55.85 | 42.40 |
| 200010 | 51.69 | 52.65 | 56.47 | 51.60 | 51.91 | 54.64 | 39.48 |
| 200011 | 50.63 | 51.80 | 55.80 | 50.23 | 51.29 | 53.93 | 40.49 |
| 200012 | 50.62 | 51.48 | 55.01 | 49.76 | 50.41 | 53.32 | 39.17 |
| 200101 | 50.22 | 51.07 | 54.66 | 48.95 | 49.89 | 52.92 | 37.95 |
| 200102 | 47.99 | 48.80 | 52.61 | 46.84 | 47.53 | 51.16 | 36.59 |
| 200103 | 44.81 | 45.68 | 50.39 | 43.56 | 45.08 | 48.25 | 36.45 |
| 200104 | 44.39 | 45.38 | 50.02 | 43.37 | 44.79 | 47.11 | 36.23 |
| 200105 | 43.24 | 44.13 | 48.90 | 41.41 | 43.14 | 45.94 | 33.44 |
| 200106 | 41.95 | 42.77 | 47.69 | 39.91 | 41.23 | 44.62 | 33.11 |
| 200107 | 41.61 | 42.71 | 46.19 | 40.07 | 41.58 | 42.90 | 32.34 |
| 200108 | 38.59 | 40.37 | 44.51 | 37.05 | 38.52 | 41.11 | 31.10 |
| 200109 | 35.79 | 37.39 | 41.32 | 33.71 | 35.33 | 39.96 | 29.92 |
| 200110 | 35.49 | 36.86 | 40.50 | 32.82 | 34.23 | 38.39 | 28.63 |
| 200111 | 33.21 | 34.85 | 39.75 | 30.19 | 32.51 | 37.22 | 28.21 |
| 200112 | 32.22 | 33.95 | 39.02 | 28.74 | 30.87 | 35.53 | 28.32 |

[^6]Appendix 18. Expenditure shares of the matched models in period $\mathbf{t - 1}$ and period $\mathbf{t}$

Table A. Televisions, refrigerators and washing machines

| $t-1$ | $t$ | Televisions |  | Refrigerators |  | Washing machines |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | $S_{M}^{t-1}$ | $s_{M}^{t}$ | $S_{M}^{t-1}$ | $S_{M}^{t}$ | $S_{M}^{t-1}$ | $S_{M}^{t}$ |
| 199902 | 199904 | 0.98 | 0.95 | 0.94 | 0.93 | 0.96 | 0.94 |
| 199904 | 199906 | 0.98 | 0.98 | 0.94 | 0.87 | 0.94 | 0.91 |
| 199906 | 199908 | 0.98 | 0.95 | 0.95 | 0.85 | 0.96 | 0.91 |
| 199908 | 199910 | 0.99 | 0.93 | 0.91 | 0.89 | 0.96 | 0.94 |
| 199910 | 199912 | 0.98 | 0.96 | 0.91 | 0.93 | 0.96 | 0.92 |
| 199912 | 200002 | 0.98 | 0.98 | 0.94 | 0.91 | 0.96 | 0.90 |
| 200002 | 200004 | 0.97 | 0.97 | 0.92 | 0.89 | 0.94 | 0.94 |
| 200004 | 200006 | 0.98 | 0.95 | 0.94 | 0.90 | 0.96 | 0.93 |
| 200006 | 200008 | 0.98 | 0.95 | 0.92 | 0.88 | 0.95 | 0.91 |
| 200008 | 200010 | 0.97 | 0.93 | 0.90 | 0.89 | 0.94 | 0.92 |
| 200010 | 200012 | 0.97 | 0.95 | 0.93 | 0.91 | 0.95 | 0.92 |
| 200012 | 200102 | 0.99 | 0.98 | 0.93 | 0.92 | 0.93 | 0.94 |
| 200102 | 200104 | 0.97 | 0.97 | 0.93 | 0.90 | 0.95 | 0.92 |
| 200104 | 200106 | 0.98 | 0.97 | 0.90 | 0.88 | 0.96 | 0.93 |
| 200106 | 200108 | 0.98 | 0.97 | 0.93 | 0.90 | 0.96 | 0.94 |
| 200108 | 200110 | 0.98 | 0.95 | 0.96 | 0.94 | 0.97 | 0.92 |
| 200110 | 200112 | 0.98 | 0.98 | 0.94 | 0.97 | 0.97 | 0.95 |
| Average |  | 0.98 | 0.96 | 0.93 | 0.90 | 0.96 | 0.93 |

Table B. Computers

| $t-1$ | $t$ | Computers |  | $t-1$ | $t$ | Computers |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | $s_{M}^{t-1}$ | $s_{M}^{t}$ |  |  | $s_{M}^{t-1}$ | $s_{M}^{t}$ |
| 199901 | 199902 | 0.83 | 0.79 | 200007 | 200008 | 0.87 | 0.78 |
| 199902 | 199903 | 0.84 | 0.79 | 200008 | 200009 | 0.79 | 0.67 |
| 199903 | 199904 | 0.72 | 0.71 | 200009 | 200010 | 0.80 | 0.85 |
| 199904 | 199905 | 0.87 | 0.76 | 200010 | 200011 | 0.88 | 0.74 |
| 199905 | 199906 | 0.86 | 0.81 | 200011 | 200012 | 0.90 | 0.81 |
| 199906 | 199907 | 0.85 | 0.80 | 200012 | 200101 | 0.81 | 0.85 |
| 199907 | 199908 | 0.89 | 0.82 | 200101 | 200102 | 0.87 | 0.83 |
| 199908 | 199909 | 0.87 | 0.70 | 200102 | 200103 | 0.88 | 0.81 |
| 199909 | 199910 | 0.85 | 0.83 | 200103 | 200104 | 0.92 | 0.88 |
| 199910 | 199911 | 0.94 | 0.70 | 200104 | 200105 | 0.88 | 0.78 |
| 199911 | 199912 | 0.88 | 0.85 | 200105 | 200106 | 0.90 | 0.82 |
| 199912 | 200001 | 0.83 | 0.82 | 200106 | 200107 | 0.84 | 0.93 |
| 200001 | 200002 | 0.75 | 0.59 | 200107 | 200108 | 0.90 | 0.76 |
| 200002 | 200003 | 0.71 | 0.61 | 200108 | 200109 | 0.88 | 0.81 |
| 200003 | 200004 | 0.85 | 0.85 | 200109 | 200110 | 0.89 | 0.81 |
| 200004 | 200005 | 0.77 | 0.71 | 200110 | 200111 | 0.87 | 0.80 |
| 200005 | 200006 | 0.86 | 0.77 | 200111 | 200112 | 0.92 | 0.76 |
| 200006 | 200007 | 0.87 | 0.82 |  |  |  |  |
|  |  |  |  | Average |  | 0.85 | 0.78 |

## Appendix 19 Average residuals for new and disappearing models

Table A. Televisions, refrigerators and washing machines

| $t-1$ | $t$ | Televisions |  | Refrigerators |  | Washing machines |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | $\bar{u}_{N}$ | $\bar{u}_{D}$ | $\bar{u}_{N}$ | $\bar{u}_{D}$ | $\bar{u}_{N}$ | $\bar{u}_{D}$ |
| 199902 | 199904 | 0.08 | 0.04 | -0.03 | -0.01 | 0.01 | 0.00 |
| 199904 | 199906 | 0.01 | 0.12 | -0.01 | 0.01 | 0.04 | -0.01 |
| 199906 | 199908 | 0.01 | 0.07 | 0.00 | -0.02 | -0.03 | -0.01 |
| 199908 | 199910 | -0.02 | 0.04 | 0.02 | -0.02 | 0.00 | 0.02 |
| 199910 | 199912 | 0.04 | 0.04 | 0.01 | 0.03 | -0.02 | -0.02 |
| 199912 | 200002 | 0.03 | 0.10 | 0.00 | 0.01 | 0.02 | -0.01 |
| 200002 | 200004 | 0.12 | -0.04 | 0.04 | -0.01 | 0.03 | 0.02 |
| 200004 | 200006 | -0.03 | 0.10 | 0.01 | -0.01 | 0.00 | -0.01 |
| 200006 | 200008 | 0.00 | 0.03 | 0.00 | 0.01 | -0.03 | 0.00 |
| 200008 | 200010 | -0.01 | 0.07 | -0.03 | -0.01 | -0.01 | 0.00 |
| 200010 | 200012 | 0.07 | 0.03 | 0.03 | -0.01 | 0.01 | -0.01 |
| 200012 | 200102 | 0.07 | 0.04 | 0.00 | -0.03 | -0.01 | 0.01 |
| 200102 | 200104 | 0.05 | -0.02 | -0.02 | 0.00 | 0.02 | 0.00 |
| 200104 | 200106 | -0.01 | 0.00 | 0.00 | -0.08 | 0.03 | -0.03 |
| 200106 | 200108 | 0.08 | 0.06 | -0.03 | 0.00 | -0.02 | 0.05 |
| 200108 | 200110 | -0.01 | 0.11 | -0.01 | -0.03 | 0.01 | 0.04 |
| 200110 | 200112 | -0.04 | 0.00 | 0.04 | -0.01 | 0.02 | -0.01 |
| Average |  | 0.02 | 0.05 | 0.00 | -0.01 | 0.00 | 0.00 |

Table B. Computers

| $t-1$ | $t$ | Computers |  | $t-1$ | $t$ | Computers |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | $\bar{u}_{N}$ | $\bar{u}_{D}$ |  |  | $\bar{u}_{N}$ | $\bar{u}_{D}$ |
| 199901 | 199902 | 0.05 | -0.02 | 200007 | 200008 | -0.03 | 0.01 |
| 199902 | 199903 | 0.01 | 0.04 | 200008 | 200009 | -0.03 | 0.00 |
| 199903 | 199904 | 0.01 | 0.05 | 200009 | 200010 | 0.00 | 0.01 |
| 199904 | 199905 | -0.02 | 0.05 | 200010 | 200011 | -0.02 | 0.01 |
| 199905 | 199906 | 0.00 | 0.03 | 200011 | 200012 | 0.05 | -0.02 |
| 199906 | 199907 | 0.03 | 0.00 | 200012 | 200101 | 0.00 | 0.01 |
| 199907 | 199908 | 0.00 | 0.09 | 200101 | 200102 | 0.01 | 0.06 |
| 199908 | 199909 | 0.01 | 0.03 | 200102 | 200103 | -0.02 | 0.14 |
| 199909 | 199910 | 0.04 | 0.05 | 200103 | 200104 | -0.05 | -0.06 |
| 199910 | 199911 | -0.01 | -0.02 | 200104 | 200105 | 0.02 | 0.06 |
| 199911 | 199912 | -0.03 | 0.05 | 200105 | 200106 | 0.01 | 0.07 |
| 199912 | 200001 | 0.02 | -0.01 | 200106 | 200107 | 0.05 | -0.11 |
| 200001 | 200002 | -0.01 | 0.03 | 200107 | 200108 | -0.11 | 0.06 |
| 200002 | 200003 | -0.01 | -0.02 | 200108 | 200109 | 0.02 | 0.04 |
| 200003 | 200004 | -0.01 | 0.01 | 200109 | 200110 | 0.05 | 0.02 |
| 200004 | 200005 | -0.01 | -0.01 | 200110 | 200111 | -0.09 | 0.17 |
| 200005 | 200006 | -0.01 | 0.04 | 200111 | 20112 | -0.01 | 0.07 |
| 200006 | 200007 | 0.03 | 0.02 |  |  |  |  |
|  |  |  |  | Average |  | 0.00 | 0.03 |


[^0]:    ${ }^{1}$ An item is defined in this paper as a durable-model, normally identified with a model number, sold in a specific type of outlet.

[^1]:    ${ }^{2}$ In this paper periods are defined by JJJJMM, where JJJJ relates to the year and MM to the month. For bimonthly periods MM relates to the second month.

[^2]:    ${ }^{3}$ The use of other types of weights in the WLS-regression does not lead to a 'standard' price index in the matched-model situation. OLS-regression leads to the Jevons price index. See Van der Grient and De Haan (2003)
    ${ }^{4}$ In section 2 it was shown that the sales of durables goods have a seasonal pattern. It is unclear whether this pattern could result in biased (chained) indexes. The new international manual on Consumer Price Indexes (ILO, 2004) does not explicitly address this topic. We have not explored this aspect in our study.

[^3]:    ${ }^{5}$ Statistics Netherlands does not publish separate indexes for televisions, refrigerators and washing machines. We computed these indexes ourselves from the official CPI data. A specific adjustment made refers to the so-called energy premium households receive when buying a lowenergy durable. This premium is not included in the scanner data prices and has therefore been excluded from the CPI.

[^4]:    ${ }^{6}$ In recent years Statistics Netherlands has been expanding its efforts to use scanner data information in improving the CPI-sample.

[^5]:    1）Coefficients in bold are significant at the $5 \%$ level in all periods

[^6]:    ${ }^{1)}$ Based on official CPI-data.

